

Patents to Products: Product Innovation and Firm Dynamics*

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Abstract

We study the relationship between patents and actual product innovation in the market, and how this relationship varies with firms' market share. We use textual analysis to create a unique data set that links patents to products in the consumer goods sector. We document that while more than half of innovation comes from never-patenting firms, patents on average reflect product innovation, but this relationship crucially depends on a firm's size. We show empirically and theoretically that as firm size increases, patent filings are less reflective of innovation in the market and are more likely to be used to deter competition.

JEL Classification Numbers: O3, O4

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1 Introduction

Product innovation – the introduction of new and improved products to the market – is a key contributor to economic growth and a central element of endogenous growth models (Romer, 1990; Aghion and Howitt, 1992). However, the paucity of detailed data about the introduction and quality of new products has led researchers to use other metrics to measure innovation. As a result, patents have emerged as the primary metric of innovation, especially after comprehensive data sets with information about patents’ timing and characteristics were made readily available (Griliches, 1981). Many great inventions such as lightbulbs and solar panels are indeed patented. Yet, no patents have been filed for other important inventions like the magnetic strip behind modern-day credit cards. In other cases, firms file patents that never turn into new products in the market.¹ While these examples suggest that the relationship between patents and product innovation is complex, we lack systematic evidence on how patents reflect actual innovation and on the underlying mechanisms governing the relationship between patents and product innovation.

In this paper we create a unique data set that links patents to products of firms in the consumer goods sector. We document that, on average, patents reflect product innovation by firms, but that this relationship largely depends on a firm’s market share. Because patents offer the legal right to exclude others from exploring same or similar ideas, firms can use patents as a protective tool to preempt competitors from entering their product market space (Gilbert and Newbery, 1982). We show empirically and theoretically that this protection is especially advantageous for large market leaders because of the incentives to defend an existing lead. Our key findings can be summarized as follows:

Fact 1: More than half of product innovation comes from firms that do not patent.

Fact 2: On average, patents are positively associated with subsequent product innovation by the patent-holding firms.

Fact 3: Larger firms have lower product innovation rates (quantity and quality), but file more patents for each new product they introduce.

Fact 4: Patenting by larger firms is strongly associated with an increase in revenue above and beyond the patents’ effect on product innovation.

Fact 5: Patenting by larger firms is associated with a decline in competing firms’ product introduction.

¹For example, in recent years, patenting activity has skyrocketed whereas innovation and productivity growth have not (Bloom, Jones, Van Reenen and Webb, 2020).

There are two main challenges in studying the relationship between patents and product innovation. First, while patent data are broadly available, measures of product innovation in the market are rarely available at large scale. To address this challenge, we use comprehensive data for firms and products sold in the consumer goods sector from 2006 to 2015 collected by Nielsen-Kilts from point-of-sale systems in retail locations. This data set includes detailed information about the characteristics of each consumer-goods product; most notably, it includes information on the product’s attributes (e.g. formula, style, content), well-measured prices, and sales. We exploit this rich data set to construct measures of product innovation. Our simplest measure is the number of new products (barcodes) introduced at the firm and product category level in a given year. Since many new products represent only minor innovations relative to existing products, we also construct measures of the quality-adjusted number of new products. We infer quality improvements by tracking the new attributes that a product brings to the market and by exploiting variation in product prices and sales.

The second challenge lies in linking product innovations to their respective patents. We address this challenge by developing two distinct matching procedures. The first procedure maps each firm’s patents to its full product portfolio using the firm names in the patent and product data sets, and allows us to construct a yearly firm-level data set (Match 1). The matching procedure in this step is simple and parsimonious, but it is too coarse. Many firms in our data are active in several product categories and could be patenting products in some product categories and not in others, so we need a more granular procedure to match patents with products. We leverage the richness of the information about product and patent characteristics in our data and use modern methods from the fields of natural language processing and information retrieval to link each firm’s patents with the products it sells (Manning, Raghavan and Schütze, 2008). For this match, we first define product categories – sets of similar products – by applying clustering analysis to the short product descriptions included in the Nielsen data extended with text from Wikipedia articles about the products. We then analyze the texts of patent applications and assign each patent to the product category with which it has the highest text-similarity score.² This categorization of firms’ products and patents results in our benchmark patent-to-products data set at the yearly firm \times product category level (Match 2).

The resulting granular data set tracks patents and products for firms in the consumer goods sector. Although our empirical results pertain to this sector, the patenting intensities and product introduction rates of these firms are, on average, comparable to those in other

²Younge and Kuhn (2016), Kelly, Papanikolaou, Seru and Taddy (2018), and Webb (2019) use similar techniques when analyzing patent documents.

manufacturing sectors. Out of 35 thousand firms covered in our data set, 15% applied for a patent at least once (9% applied during the period covered by Nielsen). This patenting rate is in line with that of the manufacturing sector and is substantially higher than that of other sectors in the economy (Graham, Grim, Islam, Marco and Miranda, 2018).

We begin our analysis by exploring the properties of patents as metrics of product innovation. First, we quantify the amount of actual innovation in the market that patents capture. Over our sample period, never-patenting firms introduced more than 54% of new products and more than 65% quality-adjusted new products. These shares are larger if we rely on the patent-to-products link at the firm \times category level. These statistics are corroborated by similar statistics about sources of growth in the sector. We decompose the 10-year sales growth of the sector into growth coming from patenting and non-patenting firm \times categories, and find that although non-patenting firms are smaller, they account for 58% of growth in the sector.

Second, we quantify how changes in patenting at the extensive margin – when firms switch to patenting – and at the intensive margin are associated with actual innovation in the market. We find that firms introduce more and better-quality products after their first patent application. Exploiting our matched firm \times category-level data over time, which allows us to control for product category-specific trends and firm-category specific effects, we find that a 10% increase in the number of patents filed a year before is associated with a 0.4% increase in product introduction. We observe similar patterns when we focus our attention on granted patents or on patents that receive many forward citations, but not when we consider non-granted patents or uncited patents. This evidence suggests that commonly used measures of the quality of patents are informative about product innovation rates.

This evidence employing unique data that directly links specific products with their underlying patents informs the literature that uses patents as proxies for innovation. The result that patenting is positively associated with product innovation is consistent with the view that patents carry the productive signal about actual innovations in the market. The strength of this association is also instructive in the context of various policies meant to encourage innovation.³ The evaluation of these policies often relies on the estimated elasticity of patents to R&D inputs. However, by and large, patents are not the main policy target of such policies – innovation is. Hence, to study how policies encouraging R&D affect product innovation, for instance, one needs to take into account not only the R&D-to-patents elasticity, but also the strength of the relationship between patents and actual innovation.

³For example, R&D tax incentives and subsidies as in Dechezleprêtre, Einiö, Martin, Nguyen and Van Reenen (2016) or Akcigit, Hanley and Stantcheva (2016a).

After documenting the properties of patents as proxies for innovation, we explore the implications of patents as a tool used by firms for protective reasons. Patents provide the legal right to exclude others from exploiting the same or similar ideas, and thus may be used to protect firms' products and deter competition. [Gilbert and Newbery \(1982\)](#) and [Blundell, Griffith and Van Reenen \(1999\)](#) suggest that market leaders have greater incentives to use preemptive patenting to protect their market lead. Survey results from [Cohen, Nelson and Walsh \(2000\)](#) report that the motives behind large firms' patenting often go beyond the direct commercialization of patented innovations and extend to the strategic deterrance of rivals. Our data allow us to provide the first direct evidence on how different firms use patents for product commercialization as opposed to using them for strategic reasons.

To uncover the protective role of patents, we start by studying how the relationship between patents and product innovation changes with a firm's market position. Using variation across firms within product categories, we estimate that firms at the bottom quintile of the size distribution in a given year, as measured by total sales in a product category, introduce one new product for every five existing products in their portfolio, on average. Firms at the top quintile of the size distribution, on the other hand, introduce one new product for every seven existing products in their portfolio. Though larger firms' innovation rates are lower, they file more patents. We show that the patent filings of larger firms have significantly weaker association with their product introduction. Moreover, the average quality improvements of new products decline more steeply with firm size than the rate of product introduction does.

The weakening relationship between patents and actual product innovation for market leaders is consistent with the idea that large firms' patents are more likely to serve a strategic purpose. We find strong evidence for this in the data. First, we find that patents filed by market leaders carry a larger sales premium, even after controlling for the quantity and quality of new products these firms introduce. By contrast, the sales premium is fully accounted by product innovation associated with patents of small firms. Second, we show that patent filings by market leaders are associated with a decline in competitors' product introduction in shared product categories. The same is not true if we consider smaller firms' patent filings. Finally, we rule out alternative hypotheses that might explain the weakening patents-to-innovation relationship, such as market leaders having more patents associated with experimentation, or longer time lags between patent filings and the introduction of products in the market, or the use of alternative strategies like licensing. In fact, we find overwhelming evidence that patents of market leaders often exhibit lower quality in terms of novelty and impact: these patents have fewer follow-up citations, more self-citations, exhibit higher textual similarity with preceding patents, and are more often used in litigation.

We build a new theoretical framework to further illustrate that, consistent with the data, market leaders have larger incentives for using their patents as a protective tool. The model builds on quality-ladder models that feature creative destruction (Aghion and Howitt, 1992), but it allows for separation between the decision to innovate and the decision to patent – a distinction we can discipline with the data set we have constructed. In the model, both innovation and patenting are costly activities. Introducing higher-quality products increases a firm’s profit, while patenting decreases the firm’s chances of being displaced by competitors. The model can replicate key empirical facts from our data. Larger firms (market leaders) shift from product innovation towards protective strategies, implying that an increase in the number of patents by large firms restricts competition and innovation and does not translate into higher consumer welfare. We use the model to provide a back-of-the-envelope calculation for the private value of a patent and, most importantly, to decompose this value into protective and productive components.⁴ The productive component represents the option value of implementing the patented idea into higher-quality products in order to gain additional profits. The protective component represents the firm’s gains from impeding creative destruction by competitors. After calibrating the model to our data, we estimate that the share of a patent’s protective versus productive value increases substantially with firm size.

Related Literature – Our findings regarding the patenting and innovation decisions of firms can speak to several puzzling macroeconomic trends in recent data: patenting is soaring, but productivity growth is stagnating (Gordon, 2016; Bloom, Jones, Van Reenen and Webb, 2020); large firms funnel more resources into intangible capital – including intellectual property, but these expenditures are manifested in the increasing dominance of those firms instead of perceptible improvements in aggregate innovation in the economy (Crouzet and Eberly, 2019); recently, Akcigit and Ates (2019, 2020) argue that the decline in knowledge diffusion from frontier firms to laggards has made an important contribution to the slowing business dynamism in the United States. Our results show that large incumbents may have limited incentives to direct their efforts towards productive rather than protective patenting, which is particularly relevant as more economic activity is reallocated towards firms with a large degree of market power (De Loecker, Eeckhout and Unger, 2020; Autor, Dorn, Katz, Patterson and Van Reenen, 2020; Gutiérrez and Philippon, 2017).

Our results contribute to our understanding of firms’ growth strategies. Recent studies

⁴This decomposition is possible because we directly observe both the sales from products linked to patents as well as the competitors’ behavior. The previous approach in the literature to infer the (total) monetary value of a patent using surveys, samples of patent sales, or patent renewals is discussed in Section 6.

have shown that large firms rely on other protective strategies such as acquiring potential competitors (Cunningham, Ma and Ederer, 2018) or forging political connections (Akcigit, Baslandze and Lotti, 2018) as their innovative activity slows down (Akcigit and Kerr, 2018; Cavenaile and Roldan, 2020). We show both theoretically and empirically that patenting is yet another protective tool that firms substitute for actual product innovation as they grow.

Given this importance of the use of protective patenting for market leaders, there is surprisingly little empirical evidence on the differential use of patents by firms with different market positions. Important papers have studied patents and the associated follow-on innovation: Williams (2013) and Sampat and Williams (2019) for human genes; Cockburn and J. MacGarvie (2011) for software products; and Lampe and Moser (2015) for follow-on patenting with patent pools. While these papers have not considered these effects by firm size, Galasso and Schankerman (2015) examined 1,357 Federal Circuit patents and showed that invalidating patent rights of large patentees led to more follow-on citations to the focal patents by small patentees. In our data, we observe direct measures of product innovation in the market for all firms in the consumer goods sector and show that patenting by market leaders is related to lower product commercialization by competitors.

Finally, our novel data set sheds light on the usefulness patent statistics for measuring innovation. In the absence of direct measures of innovation, the literature has relied on indirect inference approaches using data about employment growth (Garcia-Macia, Hsieh and Klenow, 2019) or valuing innovation from patent statistics themselves (e.g., Akcigit and Kerr, 2018). Other researchers have looked at innovations that occur outside of the patent system by examining the number of new books on technical topics (Alexopoulos, 2011) or innovations featured at World Fairs between 1851 and 1915 (Moser, 2012). While we document an overall positive association between patents and product innovation, we highlight that the usefulness of patent metrics in inferring innovation significantly depends on the market position of the firms that own the patents.

The rest of the paper is organized as follows: in Section 2, we describe the data sets and our data-matching procedures. We also discuss validation exercises and present summary statistics. In Section 4, we explore the relationship between patents and product innovation. Section 5 explores the role of firm size. Section 6 presents our theoretical framework and calculations of patent value. Section 7 concludes.

2 Patent and Product Data

2.1 Overview

We face two main challenges in our study of the relationship between patents and product innovation. First, while data about patents are broadly available, information about the introduction of new products is rarely available at large scale. Second, the link between patents and the related new products is challenging to create. This section overviews the empirical strategies we use to address these challenges.

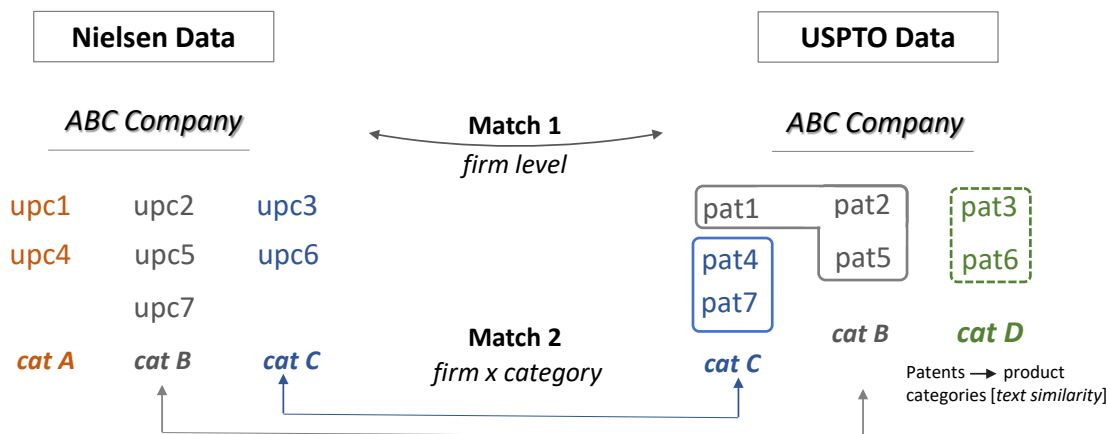
We construct a data set about product introduction beginning with product-scanner data that cover the product portfolio of firms in the consumer goods sector between 2006 and 2015. This data set allows us to identify new products by their barcodes and to observe their detailed characteristics from which we can compute various measures of innovation for a large sector. We draw patent information from the United States Patent and Trademark Office (USPTO). The combination of these two data sets gives us information about patents and product innovations covering a large sector of the economy.

To address the second challenge of linking patents to products, we develop several matching procedures. We begin by using the names of the firms in the patent and product data sets to produce a mapping between firms' patent portfolios and their respective products. We refer to this firm-level data set as Match 1. This matching procedure is simple and parsimonious, but is too coarse to allow us to connect patents with specific products. Moreover, it does not take into account that some patents are associated with process innovation, or with innovations outside the consumer goods sector.

In turn, our second matching procedure leverages the richness of product and patent characteristics using methods from the natural language processing literature to create systematic links between sets of patents and sets of products within a firm. A patent may generate no products or multiple products, and a product may have benefited from multiple patents or from none at all. Therefore, forcing a one-to-one matching between a specific narrowly defined product and a specific patent is neither possible nor desirable.

Hence, our approach is to first define product categories as sets of similar products, which are identified using clustering analysis of product descriptions extended with Wikipedia-based dictionaries. We then assign each specific patent to the product category with which it has the highest text similarity within the set of consumer goods covered by the product data. This classification of a firm's products and patents into the various product categories offered by that firm yields our benchmark patent-to-products data set, which we will refer to as Match 2. Figure 1 illustrates our data schematically, and our matching algorithms are

Figure 1: Product and Patent Data Sets



Notes: This diagram exemplifies the construction of the two data sets linking products and patents. In this example, under Match 1, all products of a firm with name “ABC Company” match to all the patents with assignee name “ABC Company”. Under Match 2, upc2, upc5, and upc7 match to pat1, pat2, and pat5 under product category B; upc3 and upc6 match to pat4 and pat7 under product category C; upc1 and upc4 of category A do not match to any patents of the firm; pat3 and pat6 of category D do not match to any products of the firm in the consumer goods sector (are either process patents or refer to products outside the consumer goods sector).

described in detail below.

To our knowledge, our algorithm generates a data set that is truly unique. [de Rassenfosse \(2018\)](#) has collected data on about 100 firms with virtual patent markings; and some private companies link patents to products of their clients.⁵ However, none of these data sets have rich systematic product-level data for all firms in a sector.

2.2 Data

Product Data – Our primary source of product information is the scanner data set from Nielsen Retail Measurement Services (RMS), provided by the Kilts-Nielsen Data Center at the University of Chicago Booth School of Business. This data set is collected from point-of-sale systems in grocery, drug, and general-merchandise stores. The original data set consists of more than one million distinct products identified by Universal Product Codes (UPCs), which are scanned at the point of sale. Each UPC consists of 12 numerical digits that are uniquely assigned to each product, and we use these to identify products. UPCs carry information about the brand and a rich set of product attributes like its size, packaging,

⁵Some examples of these companies are FairTech, IPStrategy, Powering ideas, and Intellectual Peritus. Their data sets are confidential and apply only to the portfolio of products of their clients. The service helps their clients identify their most important patents and the protection they provide to the products in their portfolio as well as helps firms prepare in case of prosecution. These companies also use text similarities techniques, using various parts of the patents text and short descriptions of products from trademark data and other sources.

formula, and flavor.

The data focus on the consumer product goods (CPG) sector, which accounts for 14% of the total consumption of goods in the U.S. This sector includes food and non-food categories (health and beauty aids, non-food grocery, and general merchandise such as cookware, electronics, gardening, household supplies). Our data cover the years from 2006-2015, and combines all sales, quantities, and prices at the national and annual levels. We use the panel structure of each UPC to measure its entry year. This product data set covers about 40% of the CPG sector sales, and nearly the universe of firms and new products in the sector. Appendix [A.I](#) provides additional details about the coverage and representativeness of Nielsen RMS to measure product innovation in the consumer goods sector.

Patent Data – Our main source of data for patent analysis is the USPTO data on the universe of published patent applications, granted or not. We use the original bulk data files provided by USPTO’s Bulk Data Storage System for our analysis. Our sample initially contains information on more than 7 million patent applications filed by more than 500 thousand patent assignees in the years 1975-2017. For each patent, we use information about the patent application year, patent status (granted, pending, or abandoned), patent technology classifications, forward patent citations received, the number of claims on a patent, and whether it is a utility or design patent. For our textual analysis of patent documents, we extract patent titles, the text of patent abstracts, the text of corresponding patent classification titles, claims text, and the titles of citing patents. Appendix [A.II](#) gives more detail about our sample and the variables we use.

2.3 Matching Firms

In our firm-level data set (Match 1), we match patents to products at the firm level using the firm names in both patent and product data sets. To match firms to patents, we obtain the firm names for each product using data from the GS1 US, which is the single official source of UPCs. This data set links barcodes with the names of firms that sell the product. For the patent data, we begin with the assignee name(s) of each patent. This name is typically the original assignee of the patent and may not represent the current user of the patent because of sales or company reorganizations. We combine the USPTO patent re-assignment data with Thomson Reuters Mergers & Acquisition data to re-assign each patent to its most current holder. This step relies on the assumption that when a firm acquires (or merges with) another firm, the new firm will own all patents that the firms owned before the acquisition (merger). The details of these steps are described in Appendix [A.I](#) and [A.II](#).

A firm’s name could be formatted or abbreviated differently in the product and patent data sets, or it may even be misspelled, which presents a challenge in joining information from the two data sets. We developed a name-cleaning algorithm to clean and standardize the firm names to overcome this challenge. This procedure builds on and extends cleaning algorithms from the NBER Patent Data Project (Hall et al., 2001) and Akcigit et al. (2016b) and is described in detail in Appendix A.III.

2.4 Matching Patents to Product Categories

The algorithm used to build Match 2 (Firm \times Category level) has three crucial steps. In this section, we describe the details of these steps. The first step creates product categories at a level of aggregation such that they collect distinct and sufficiently large sets of similar products that are meaningfully related to a distinct set of patents. This step yields a set of product categories, a vector of terms used to describe each product category, and a mapping of products into categories. In the second step, similarity scores between patents and product categories are computed. We use various text descriptions to build a vector of terms that describe each patent. We then compute similarity scores between each patent and every product category. These scores represent the overlap between the texts in patents and the text associated with each product category. The final step of our patent-product matching algorithm consists in using the similarity scores and information about the production of the respective patenting firms to classify each patent into a product category and filter out patents not related to CPG products.

Defining Product Categories – We define product categories by exploring the product classification scheme used by Nielsen. In the original data, each product is classified into one of 1,070 detailed product modules. These product modules are further aggregated into a set of 114 product groups, and those are further aggregated into ten departments. For example, “disposable cups” and “disposable dishes” are two distinct product modules that are part of the group “paper products” which is part of the department “non-food grocery”. Nielsen’s modules aggregate products that are close in their technological characteristics. However, there are some sets of distinct modules that have very similar products. At the same time, many Nielsen’s groups include products that are quite distinct. For example, “disposable cups”, “disposable dishes”, “pre-moistened towelettes” and “paper napkins” are all part of the group “paper products”, but only “disposable cups” and “disposable dishes” are technically similar. Hence, we seek an intermediate categorization of products – more aggregated than modules and less aggregated than groups – to be able to meaningfully

associate patents to a well-defined set of products.

To this end, we apply a clustering procedure to aggregate the Nielsen modules into distinct product categories. Each module is paired with a vector of descriptive terms (single words and two-word phrases) that are weighted by their importance. We expand short module descriptions from Nielsen data with the text of hand-collected Wikipedia articles to get to a comprehensive description of the product content of the modules. The resulting vectors of descriptive terms collect all the words from the Wikipedia and Nielsen texts, after applying standard parsing and lemmatizing algorithms. When building term vectors, one must appropriately weight terms by their importance. We use the leading approach in textual analysis – the “term-frequency-inverse-document-frequency” sublinear transformation – that accounts for both the frequency with which a term appears describing a module and how commonly it is used to describe other modules (Aizawa, 2003). This approach ensures that we under-weight common terms that appear in many documents as these are less diagnostic of the content of any individual document.

We then aggregate these module vectors into clusters using a popular technique known as k-means clustering (Lloyd, 1982). This procedure allows one to specify the desired number of clusters and yields a clustering assignment that minimizes the within-cluster term vector variance. As a baseline, we use an aggregation of modules into 400 clusters which we refer to as product categories. We find that this partition strikes a balance between aggregating very similar products while maximizing the difference between products across categories. Appendices A.IV.1 and A.IV.2 provide extensive descriptions of methods we have taken from the literature on natural language processing, including the details of clustering, quality assessment, and alternative methods to encourage robustness. Appendix F shows the sensitivity of our main findings using the aggregation of product modules defined by Nielsen.

After defining the level of aggregation, we build term vectors describing each product category. We use the same methodology that we used to build the term vectors for modules, but now we use the titles of the clustered module(s) and all the text from their corresponding Wikipedia articles. We ensure that when a product category aggregates multiple modules, we first vectorize each module description and then average these vectors together so that we do not overweight longer entries. The final product category vectors are normalized to have unit length.

Patent Vectors and Similarity Scores – This subsection describes how we measure the amount of overlap between the texts of patent applications and product categories. For the patent description, we use the following fields from the patent applications: the title,

abstract, international patent classification description, and the titles of cited patents. We create vectors of terms by concatenating all these fields into one document, followed by the same parsing and lemmatizing algorithms. As before, we adjust the weights of each term according to the “term-frequency-inverse-document-frequency” sublinear transformation and normalize patent vectors to have unit length.

Finally, we construct a similarity score for each patent p and each product category j by computing the cosine similarity between two normalized vectors, $s_{jp} = f_j \times f_p$. This similarity score is guaranteed to be in the range $[0, 1]$ with zero indicating no word overlap and one indicating that the documents are identical. Appendix A.IV.3 provides technical description of this step.

Classifying Patents into Product Categories – The final step of our patent-products matching algorithm consists in using the similarity scores and determining which product categories are valid matches for each patent. We must, however, make some adjustments because we use all patents of each firm with products in the consumer goods sector, and some patents may relate to goods outside the consumer goods sector or correspond to more general process/method patents. Hence, we should allow for the possibility that a patent will not be assigned to any product category. After an extensive review of patent texts and a great deal of testing, we identified systematic adjustments to the algorithm that ensure that irrelevant patents remain unmatched with products.

We first adjust the algorithm to include a similarity score threshold. We tested different threshold levels and, in our baseline algorithm, we restrict the set of potential categories for each patent p to the product categories whose similarity score exceeds 0.025. The idea is that patents with low text similarity are unrelated to the product categories that we consider. The implication of this adjustment is that patents whose highest similarity are below that threshold are more likely to be classified as “non-matched”.

Second, we use information about the set of product categories sold by the firm. For each patent, we define the set of potential matches, whose elements consist of all product categories in which the patenting firm ever sold a product, according to our product data. Together, these criteria imply that patent p will be classified as unmatched if no product categories satisfy the threshold similarity and belong to the set of categories the firm produces in. For the patents that have more than one product category satisfying those conditions, we assign the final patent-product category match so that the patent matches to the product category with the highest similarity score.

Our methodology assumes one product category match for each patent. However, some patents may be more general in nature so that they relate to multiple categories. Our

baseline algorithm abstracts from this possibility. However, our procedure to define product categories is designed to ensure that the product categories would encompass a broad range of products that are technically similar such that one patent plausibly relates to this and only this range of products.⁶ In Appendix A.IV.4, we present the details of this procedure and all the robustness exercises with which we tested our baseline algorithm. Appendix F shows the sensitivity of our main findings under higher similarity thresholds.

2.5 Match Statistics and Validation

Table 1 provides statistics of the baseline data used in our analysis. The data set Match 1 (Firm level) includes annual data for all 34,665 firms that sold at least one product in our consumer goods sector data (CPG firms). The raw USPTO patent data cover information from 1975 to 2017, but because our product data only cover years from 2006 to 2015, our analysis can only consider annual variation for the period 2006-2015. In this shorter period, the USPTO data include about 3.4 million patent applications in total, and about 500 thousand patent applications filed by CPG firms. The data set Match 2 (Firm \times Category level) includes 40% of those patent applications.⁷ The remaining 60% of patents, while filed by CPG firms, could not be associated with products in the consumer goods sector.

We perform an extensive set of validation exercises to evaluate the robustness and quality of our match. Appendix A.V presents details on these validation exercises, while here we focus on summarizing the most important. We use four main types of validation exercises: manual checks, external validations using online-collected data on patent markings, analysis of the robustness of the algorithm-implied similarity scores and placebo tests, and validation of non-matches.

Manual checks – We manually checked many of the patent-to-products matches and some examples are listed in Table A.I in the Appendix. The table lists 100 patent applications by the top-selling firms in the largest product categories according to Nielsen. One can easily see that the patent titles reflect well the product categories to which the patents were assigned. For most patents we analyzed, we found that our manual choices of product categories also coincide with the product categories chosen by our matching algorithm using similarity scores.

⁶In this sense, the methodology delivers a many-to-many patent-to-product match, where each patent can be matched to multiple products of the firm.

⁷Appendix B illustrates examples of a variety of patent applications and the corresponding products with their respective product entry dates.

Table 1: Match Statistics

| | Period | |
|-------------------------------------------------|---------------|-----------|
| | 1975-2017 | 2006-2015 |
| Number of patent applications | | |
| All assignees in USPTO | 7,304,072 | 3,386,208 |
| CPG firms (Match 1) | 1,046,030 | 505,544 |
| CPG firms in product categories (Match 2) | 399,684 | 190,575 |
| Number of firms | | |
| All CPG firms | | 34,665 |
| CPG with at least a patent applied in 1975-2017 | | 5,209 |
| CPG with a patent applied in 2006-2015 | | 3,266 |

Notes: Match statistics for the baseline data sets Match 1 (Firm level) and Match 2 (Firm \times Category level). Match 1 is described in Section 2.3 and Match 2 is described in Section 2.4.

Virtual patent markings – We next use virtual patent markings to validate our matches. Using virtual patent markings, firms may give a notice to the public that their product is patented by publishing their products and the patents protecting them online. Website searches showed that very few firms in our data used virtual patent markings, and even when they did, only a selection of products and patents appeared in the markings. Nevertheless, these data give a unique opportunity for an external validation of our matching algorithm.

For Procter & Gamble (P&G) and Kimberly Clark (KC), we manually collected virtual patent markings from the company websites and mapped them to our product categorization. We then validate our patent-product category matches for these firms against this information. Appendix A.V.2 shows that the patent-product category mapping from virtual markings is also selected by our matching algorithm in about 70% of cases.

Robustness of the match and placebo tests – We evaluate robustness of the product category choice by our matching algorithm to potential small perturbations in the algorithm. For the algorithm to be robust against small changes, we should observe that highest-ranked product categories have substantially higher similarity scores with the patents than lower-rank product categories do. Section A.V.3 in the Appendix shows this is the case. Next we verify that we are indeed carving out well-defined neighborhoods in the technological space by matching patents into distinct categories. For that, we compare the actual distribution of similarity scores between patents classified in the same product category versus a placebo group of patents drawn at random. Section A.V.4 in the Appendix shows that the distribution of similarity scores between pairs of patents within product categories is indeed very different and first order stochastically dominates that of the placebo group.

Validating non-matches – In our last step of the algorithm for Match 2, multiple criteria are used to allow for the possibility that some patents filed by CPG firms are not associated with any of the consumer-good product categories. A valid “non-match” can arise for two main reasons. First, a patent may relate to goods that the firm may be producing outside the CPG sector; second, a patent may be about a general process or method that does not affect the introduction of new products. In the spirit of [Hoberg and Phillips \(2016\)](#), we use information from publicly traded companies’ 10K reports to identify firms whose output is mostly in the consumer-goods sector, and we find that only a minority of their patents are classified as “non-match”, contrasting with patents held by firms who mostly sell products outside the consumer goods sector. Next, we use alternative procedures to proxy for process patents (completely independently from the algorithm) and compare them with the algorithm’s “non-matches”. We follow [Bena and Simintzi \(2017\)](#), and use patent claims to create proxies for process-related and product-related patents. We find that the share of “non-matches” is significantly higher among the claims-based measure of process-related patents. These exercises, which are presented in Section [A.V.5](#) in the Appendix, offer reassurance that our algorithm successfully filters out patents that are not directly related to the products in our data.

3 Measures of Product Innovation and Patenting

3.1 Product Innovation

Our measures of product innovation are based on the number of products that firms introduce to the market and the quality improvements in those products. We use the product data described above to identify the entry dates of products in the market and their respective characteristics and performance. We create separate measures of innovation for the firm-level (Match 1) and firm×category level (Match 2) data. Our first measure is the number of **new products** of firm i (in product category j) in year t , as in [Broda and Weinstein \(2010\)](#), [Argente, Lee and Moreira \(2018\)](#) and [Jaravel \(2019\)](#):

$$N_{ijt} \equiv \sum_{u=1}^{T_{ijt}} \mathbb{1}[u \text{ is entrant}],$$

where product u is sold by firm i in product category j , T_{ijt} is the number of products that firm i sells in j as of period t , and $\mathbb{1}[u \text{ is entrant}]$ is an indicator that takes the value of one if u is a new barcode in t . This measure is simple and parsimonious but does not distinguish major product innovations from innovations that make relatively minor changes

to a product’s characteristics. In contrast to the previous literature, we construct the second set of measures of **quality-adjusted new products** that deals with this potential drawback by explicitly accounting for differences in characteristics across new products:

$$qN_{ijt} \equiv \sum_{u=1}^{T_{ijt}} q_u \mathbb{1}[u \text{ is entrant}],$$

where $q_u \in [0, 1]$ is a measure of quality that we describe below. Together these two metrics allow us to account for differences in both the quantity and quality of product innovation across firms and over time.

Our baseline measure of product quality aims at capturing differences in novelty and economic impact across new products. We build on [Argente and Yeh \(2017\)](#) and use detailed information on product attributes that is available from the product data. Products can then be compared on the basis of characteristics associated with their attributes $\{v_{u,1}, \dots, v_{u,A}\}$.⁸ We test if each new product has characteristics distinct from those of all existing products available in the market, and we compute the quality of a new product as a weighted sum of its novel characteristics across all product attributes:

$$q_u \equiv \sum_{a=1}^A \omega_a \mathbb{1}[v_{ua} \text{ is new}].$$

where ω_a are weights that reflect the economic value associated with a particular attribute. We develop a novel approach to estimate weights that capture the importance of each attribute by using “shadow prices” from hedonic pricing regressions ([Bresnahan and Gordon, 1996](#)). The underlying assumptions here are that the degree of novelty of a product should be reflected in the price of a product and that the price of a product reflects its embodied characteristics as valued by shadow prices. A new product has a high novelty score if it has many new characteristics and/or if its characteristics are associated with high implicit prices. We provide details on the properties of this procedure in [Appendix C](#), along with some evidence that the novelty score is strongly associated with the performance of the firm and its products.⁹

⁸For example, “children” and “regular” are two mutually exclusive characteristics associated with the attribute “formula” for “pain remedies-headache” products. Naturally, the number and type of attributes varies across product categories. For example, the product category “pain remedies-headache” includes 10 attributes: brand, flavor, container, style (i.e. children, regular), form, generic, formula (i.e. regular, extra strength, rapid release), type (i.e. aspirin), consumer (i.e. trauma, migraine), and size. On average, we observe that the different product categories include between 5 to 12 attributes. [Appendix C](#) gives details.

⁹We show that our measure is correlated with the growth rate of the firm, the share of sales generated by new products, and the average duration of new products in the market even after conditioning on the number of products being introduced by the firm ([Table A.II](#) in the Appendix).

We use three alternative measures of new product quality to evaluate the robustness of our empirical results. First, we use a simpler version of the quality measure that weighs each attribute equally (quality $q1$). This measure only captures variation in the share of new product characteristics contained in a product. Second, we use a weighted quality measure using weights that reflect “shadow sales” (quality $q2$). This measure assigns lower quality to new products that are associated with high shadow prices but do not reach many customers. Finally, we use a measure of residual demand taken from [Hottman et al. \(2016\)](#) and [Argente et al. \(2020\)](#) (quality $q3$). This measure does not use information about the degree of novelty of a product and instead captures the relative appeal of new products relative to other products sold in the market, under some functional-form assumptions. Overall, our baseline measure and these alternative metrics allow us to consider many critical dimensions of the quality of new products, and allow us to assess the robustness of our results.

3.2 Patent Measures

Using an approach similar to how we measured product innovation, we compute measures that allow us to account for differences in the quantity and quality of patent applications across firms and over time. Our baseline measure is the number of **patent applications** (P_{it}). Using our patent-product category match, we are also able to measure the number of patent applications filed by firm i in product category j in year t as follows:

$$P_{ijt} \equiv \sum_{p=1}^{P_{it}} \mathbb{1}[p \text{ is match to } j].$$

Throughout the paper, we use information about whether a patent was granted and information about patent citation counts to compute our measures of patent quality. Patent applications that become **granted patents** (gP_{ijt}) are perceived as high-quality patents because the patent office deemed them novel enough to not be rejected. We compute the number of patent applications that are granted as:¹⁰

$$gP_{ijt} \equiv \sum_{p=1}^{P_{it}} \mathbb{1}[p \text{ is granted}] \times \mathbb{1}[p \text{ is match to } j].$$

We also define **patent citations** (cP_{ijt}) as the total number of patents weighted by forward citations received in the first five years since the application was filed:¹¹

¹⁰The condition $\mathbb{1}[p \text{ is match to } j]$ is only used for Match 2.

¹¹A 5-year citations measure attempts to reduce the truncation issue inherent to citations – the fact that patents filed more recently have had less time to accumulate citations ([Hall et al., 2001](#)).

$$cP_{ijt} \equiv \sum_{p=1}^{P_{it}} c_p \times \mathbb{1}[p \text{ is match to } j].$$

Measures based on forward citations have traditionally been used to assess the economic and technological significance of a patent (for earlier contributions, see [Pakes \(1986\)](#), [Schankerman and Pakes \(1986\)](#), [Trajtenberg \(1990\)](#)).

3.3 Summary Statistics

Table 2 provides summary statistics about the product- and patent-related variables for the firms in our sample, grouped by their patenting activity. We split firms into three groups: (i) firms that have never filed a patent application, (ii) firms whose last patent application was filed before 2006 (the beginning of the Nielsen RMS data set) and (iii) firms that filed a patent application between 2006 and 2015.

The share of patenting firms and product introduction rates in the consumer goods sector are comparable to those of other manufacturing sectors. Table 2 shows that more than 5 thousand firms (15%) applied for at least one patent and more than 3 thousand firms (9.5%) filed a patent application during the period 2006-2015. For comparison, [Graham et al. \(2018\)](#) links Census data to the USPTO and finds that 6.3% of manufacturing firms have at least one granted patent application between 2000 and 2011.¹² The corresponding number in our data is 7.6%, which is only slightly higher.¹³ Table 2 indicates that product introduction rates are on average 20%. While there is no equivalent comprehensive product data for other sectors, [Goolsbee and Klenow \(2018\)](#) use the Adobe Analytics data on online transactions covering multiple products and report product introduction rates that are comparable to those of other non-durable consumer manufacturing sectors.¹⁴

As expected, patenting firms are larger: they sell more products, operate in more product categories, and have higher sales. Firms that filed patents between 2006 and 2015 account for 61% of sales in our sample. Patenting firms also introduce more products, but this relationship is weaker once we account for scale and instead focus on the rates with which new products are introduced. Interestingly, our three different quality measures indicate that the average novelty of new products sold by patenting firms is not higher than that of non-patenting firms, conditional on product introduction.

Firms with patent applications between 2006 and 2015 file more than six patents per

¹²The incidence of patenting in the rest of the economy is lower, at 1%.

¹³Notice that [Graham et al. \(2018\)](#)'s patent data include only granted patents, while our data also include unsuccessful patent applications. If we count only granted applications, we would have 2629 patenting firms.

¹⁴[Goolsbee and Klenow \(2018\)](#) show that some durable consumer goods (e.g. furniture), not covered in our data set, have entry rates that are larger than those of non-durables (e.g. food).

Table 2: Summary Statistics by Firm’s Patenting Status

| | No Patents | Patents before 2006 | Patents 2006-2015 |
|-----------------------------------------------------------|------------|------------------------|----------------------|
| Product data | | | |
| Number of products | 15.49 | 31.08 | 78.35 |
| Number of new products (N) | 2.58 | 5.26 | 13.45 |
| Average quality of new products (q) | 0.27 | 0.20 | 0.20 |
| Quality-adjusted number of new products (qN) | 0.46 | 0.62 | 1.48 |
| Product introduction rate (n) | 0.19 | 0.17 | 0.22 |
| Quality-adjusted product introduction rate (qn) | 0.07 | 0.04 | 0.06 |
| Sales from all products | 2371.59 | 9392.09 | 37094.71 |
| Sales from new products | 454.74 | 1811.01 | 8130.00 |
| Number of product categories | 2.36 | 3.07 | 5.46 |
| Average quality of new products ($q1$) | 0.13 | 0.10 | 0.10 |
| Average quality of new products ($q2$) | 0.18 | 0.11 | 0.12 |
| Average quality of new products ($q3$) | 0.06 | 0.32 | 0.10 |
| Patent data | | | |
| Number of patent applications (P) | 0.00 | 0.00 | 6.34 |
| Number of granted patent applications (gP) | 0.00 | 0.00 | 4.57 |
| Number of citations-weighted patent applications (cP) | 0.00 | 0.00 | 5.88 |
| Stock of patent applications | 0.00 | 11.33 | 125.36 |
| Stock of granted patent applications | 0.00 | 11.02 | 107.63 |
| Stock of citations-weighted patent applications | 0.00 | 17.97 | 215.24 |
| Number of firms | 29215 | 1943 | 3266 |
| Observations | 186934 | 15803 | 29052 |

Notes: The table shows the average of product-based and patent-based variables of the Match 1 data set. The first column groups firms that have no patents; the second column considers firms that have patents, but filed them before they first appear in Nielsen RMS (before 2006); and the third column is for firms that have patents in our focus period of 2006-2015. The statistics regarding product introduction can only be computed for the period 2007-2015 because we cannot determine entries for products first introduced in 2006 (left censored). The statistics for sales are given in thousands of dollars, deflated by the Consumer Price Index for all urban consumers. Patent statistics are very skewed; the table reports averages after winsorizing patent-based variables at top 0.1%.

year, on average. Because many patents receive no citations, especially in the first five years, the average number of citation-weighted patent applications, cP_{ijt} , is very similar to the average raw number of patent applications, P_{it} . These firms may hold some design patents, but the majority of patents in our sample are utility patents. Unsurprisingly, the summary statistics show that firms who filed a patent between 2006 and 2015 hold a larger stock of patents than firms who last filed a patent application before 2006.

Our data cover product categories that exhibit substantial heterogeneity in entry rates and patenting intensity. In Appendix E we provide some descriptive statistics grouped across food and non-food categories.¹⁵ The two types of product categories have, on average,

¹⁵Throughout the analysis we mostly use variation within detailed product categories which do not capture heterogeneity in innovation and patenting intensities across types of products. Nevertheless, in the Appendix E we provide the results using only food or non-food product categories to ensure that the results are not driven by variation within some specific product categories.

similar entry rates but distinct patent intensities. The share of patenting firms and the ratio of patents per product is higher for non-food categories such as health and beauty aids (including over-the-counter drugs), non-food grocery, and general merchandise (including cookware, electronics, various household supplies).

4 Relationship Between Product Innovation and Patents

In this section, we explore the properties of patents as metrics of product innovation by evaluating how patents relate to actual product introduction in the market, and how much product innovation is captured by patent-based metrics of innovation. First, we show the cross-sectional allocation of product innovation between patenting and non-patenting firms. Second, we consider how a firm’s product introduction changes after it files a first patent application. Third, we quantify the strength of the relationship between the changes in the number of patents filed and the amount of product innovation. Finally, we explore the dynamics of these relationships. Our findings can be summarized in two empirical facts:

Fact 1: More than half of product innovation comes from firms that do not patent.

Fact 2: On average, patents are positively associated with subsequent product innovation by firms.

Product Introduction Accounted by Patenting Firms – We begin our analysis of the relationship between patents and product innovation by exploring cross-sectional variation across firms according to their patenting status. Table 3 shows that in our data, 54% of new products were introduced by firms that never applied for a patent. If we account for the degree of novelty of new products, we estimate that about 65% of quality-adjusted product introduction comes from never-patenting firms.¹⁶ This indicates that, on average, patenting firms introduce more products that make only an incremental improvement over existing products on the market.¹⁷

Since they rely on the firm-level match, the above statistics implicitly attribute all new products introduced by a patenting firm to some of its patents. However, highly diversified firms might be patenting in one product category, while introducing many products that have no relation to the patents they are filing in other categories. Thus we may be attributing

¹⁶Note that our calculations only speak to the *direct* contribution of patents in product innovation and would not account for indirect channels such as between-firm knowledge spillovers.

¹⁷This observation holds true regardless of the quality adjustment we use. For example, the share of $q1N$ accounted by never-patenting firms is 65%, and the share of $q2N$ by never-patenting firms is 77%. Our residual quality measure of innovation, $q3$, does not allow us to construct a good counterpart to $q3N$, however as seen from Table 2, $q3$ is not necessarily higher for patenting firms.

Table 3: Share of Product Innovation Accounted for by Patenting Firms

| | New Products, N | Quality-adjusted New Products, qN |
|--------------------------------------------------|-------------------|----------------------------------------|
| Match 1 | | |
| Firms with patents in 2006-2015 | 0.38 | 0.28 |
| Firms with patents before 2006 | 0.08 | 0.07 |
| Firms with no patents | 0.54 | 0.65 |
| Match 2 | | |
| Firm \times category with patents in 2006-2015 | 0.23 | 0.16 |
| Firm \times category with patents before 2006 | 0.07 | 0.05 |
| Firm \times category with no patents | 0.71 | 0.79 |

Notes: the table shows the share of product innovation in the market measured by our two benchmark measures – product introduction (column 1) and quality-adjusted product introduction (column 2) – accounted for by firms and firm \times categories with or without patents.

too much product introduction to patents if we rely only on the firm’s overall patenting status. This observation exemplifies the importance of establishing a closer link between patents and products using the Match 2 data set. To make these more granular links, we replicate the above exercise but define patenting status at the firm \times category level. As seen from Table 3, firms that never patented in a category are responsible for a greater share of new products introduced in that category.

It is not surprising that a large fraction of innovation may not be directly associated with specific patents. Even if firms wanted to patent all new products, some products represent only small upgrades to existing ones and thus may not be patentable. Patents are only granted if they exhibit “novelty and non-obviousness,” and new products that result from incremental changes will likely not be captured by raw patent metrics. So while it is natural that some innovations are not captured by patent statistics, our data offer a unique opportunity to quantify the magnitude of this omission. There is a dearth of empirical evidence on the extent to which new market innovations are reflected in patent statistics, so we believe this exercise is especially useful given their widespread use.¹⁸

We also evaluate if our measures of product innovation reflect well the sources of growth. Indeed, if we look through the lens of classic innovation-driven growth models, we should expect innovation and growth measures to go hand-in-hand. We conduct simple growth decompositions for our sector to get at this question. We decompose sales growth from 2006 to 2015 into growth that comes from patenting and non-patenting firm \times categories as:

¹⁸The only exceptions we are aware of are the studies by Moser (2012) and Sampat and Williams (2019). The first study identifies all innovations featured at World’s Fairs between 1851 and 1915 and documents those that were patented; while the second analyzes the full sample of human genes and identifies those whose sequences were claimed in U.S. patents.

$$\underbrace{\text{Growth}_{06-15}}_{7\%} = \underbrace{\text{Growth}_{06-15}^{\text{Patent}}}_{4\%} \times \underbrace{s_{2006}^{\text{Patent}}}_{0.72} + \underbrace{\text{Growth}_{06-15}^{\text{No Patent}}}_{14.4\%} \times \underbrace{s_{2006}^{\text{No Patent}}}_{0.28} \quad (1)$$

where s_{2006}^{Patent} and $s_{2006}^{\text{No Patent}}$ denote sales shares of firm \times categories with or without patents, respectively.¹⁹ As with our measures of product innovation, these growth decompositions show that although non-patenting firms are smaller and account for a smaller share of sales in the sector, they contribute more to growth relative to the set of patenting firms – totaling to 58% of the sectoral growth. Hence, the fact that more than half of the product innovation in the sector is not captured by the patenting status of the firms is corroborated by similar statistics about growth.

First-time Patent Filers – One important feature of our data is that we observe some firms that change their patenting status in the period of analysis 2006–2015. This allows us to evaluate whether a firm’s product introduction tends to change after the firm’s first patent application.²⁰ We do so by estimating the following specification:

$$\log Y_{it} = \beta dP_{it} + \alpha_i + \gamma_t + u_{it} \quad (2)$$

where Y_{it} is the outcome of firm i in year t , α_i represents firm fixed effects, and γ_t represents year effects. dP_{it} is an indicator variable that equals 1 after the firm’s first patent application. Our goal is to understand if the switch to patenting is associated with increased product innovation, which would be the case if patent-based measures were to approximate well product innovation in the market. To uncover this relationship, we estimate the effects of β relative to firms that are already patenting. These firms have more similar characteristics and thus are likely a more suitable counterfactual for firms that first apply for a patent than those that never apply.²¹

Table 4 presents the estimated change in our two measures of product innovation associated with a firm’s transition from non-patenting to patenting. Conditional on firm and year effects, we find an average increase in product introduction of up to 11% after the switch

¹⁹We first write $Rev_t^{CPG} = \sum_j \sum_{i \in \Omega_{\text{Patent}}^j} Rev_{ijt} + \sum_h \sum_{i \in \Omega_{\text{No Patent}}^j} Rev_{ijt}$, where the second sum is across product categories and Ω denotes the set of firms with and without patents in category j ; and take the percentage changes in sales to arrive at (1).

²⁰Although defining the event of the first patent application at the firm level (as opposed to firm-category) avoids potential cross-category spillovers in patenting and sharpens our definition of the event, we also explore the dynamics of this relationship in detail using Match 2 data later on.

²¹The assumption that this group of firms forms a better control – after accounting for time-invariant differences between firms and common year factors – is supported by the summary statistics presented in Table 2. Nevertheless, we find similar estimates when we test if our results are explained by the contrast with the entire sample of non-switching firms, which includes firms already patenting before the beginning of our sample and those that have not yet patented at the end of our sample (Table A.III in the Appendix).

Table 4: Product Innovation after First Patent Application

| | Log N | | | Log qN | | |
|-----------------------------|---------------------|---------------------|--------------------|-------------------|---------------------|--------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| After patent(t) | 0.1168** (0.045) | | | 0.0352 (0.020) | | |
| After granted patent(t) | | 0.1361** (0.048) | | | 0.0497** (0.018) | |
| After non-granted patent(t) | | | -0.0045 (0.044) | | | -0.0085 (0.036) |
| Observations | 29,470 | 29,470 | 29,470 | 29,470 | 29,470 | 29,470 |
| Time | Y | Y | Y | Y | Y | Y |
| Firm | Y | Y | Y | Y | Y | Y |

Notes: The table shows regressions of log number of new products (Log N) in Panel A and of log quality-adjusted new products (Log qN) of a firm as a function of a dummy equal to one after the first patent application by the firm. Our benchmark quality measure is defined in Section 3.1. The alternative innovation quality measures ($q1, q2, q3$) produce similar results. Both Log N and Log qN use the inverse hyperbolic sine transformation. *After patent* is a dummy equal to one after any patent application; *After granted patent* is a dummy equal to one after a patent application that is granted; and *After non-granted patent* is a dummy equal to one after a patent application that has not been granted (abandoned or pending). The sample includes 596 firms that switch to patenting and those that already patented before the beginning of our sample. Standard errors robust against heteroskedasticity and serial correlation are reported in parentheses.

to patenting. Columns (2) and (3) show that the positive correlation is largely driven by high-quality patent applications, if we take the patent’s success with the patent office as a proxy of quality. This result is more pronounced if we study the effect on quality-adjusted product introduction, as shown in columns (4) to (6). These exercises reveal a positive correlation between the timing of patent applications and product innovation.

Intensive Margin of Patenting – We next explore how product innovation varies with the changes in intensive margin of patenting exploiting variation in measures of product innovation at the firm \times category level over time. We estimate

$$\log Y_{ijt} = \beta \log P_{ijt-1} + \alpha_{ij} + \gamma_{jt} + u_{ijt} \quad (3)$$

where Y_{ijt} is the outcome for firm i in category j in year t and P_{ijt-1} is the log number of patent applications filed by the firm i in category j a year before to allow for a short lag between patent filing and product commercialization. Thanks to the firm \times category level data, we can now control for product category-specific trends (e.g., market-wide demand for specific products), and we can control for firm-category specific effects, thus filtering out, for instance, the effects of firm-specific market power on the sales of specific products. Importantly, this set of fixed effects also ensures results are not driven by differences in patentability or coverage across distinct product categories, or firm specific time-invariant predispositions to apply for patents.

Table 5: Product Innovation and Patenting

| | Log N | | | Log qN | | |
|--------------------------|----------------------|----------------------|--------------------|----------------------|----------------------|-------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Patents(t-1) | 0.0380*** (0.009) | | | 0.0189*** (0.005) | | |
| Patents granted(t-1) | | 0.0405*** (0.010) | | | 0.0192*** (0.005) | |
| Patents non-granted(t-1) | | | 0.0234* (0.013) | | | 0.0082 (0.007) |
| Observations | 409,641 | 409,641 | 409,641 | 409,641 | 409,641 | 409,641 |
| R-squared | 0.692 | 0.692 | 0.692 | 0.623 | 0.623 | 0.623 |
| Time-Category | Y | Y | Y | Y | Y | Y |
| Firm-Category | Y | Y | Y | Y | Y | Y |

Notes: The table shows regressions of the log number of new products ($\log qN$) and of log quality-adjusted new products ($\log qN$) in a firm \times category over time as a function of the log number of patents. Our benchmark quality measure is defined in Section 3.1. The alternative innovation-quality measures ($q1, q2, q3$) produce consistent results. *Patents* is the log number of any patent applications in firm \times category \times year; *Patents granted* is the log number of granted patent applications; and *Patents non-granted* is the log number of patent application that have not been granted (abandoned or pending). The inverse hyperbolic sine transformation is used for logs. Standard errors robust against heteroskedasticity and serial correlation are reported in parentheses.

Table 5 shows the estimates. The rows present results from using different explanatory variables – the log number of patents, granted patents, and non-granted patents. Conditional on firm-category and category-time fixed effects, we find that the observed elasticities of product introduction and quality-adjusted product introduction to patents are 0.04 and 0.02, respectively. As before, the relationship between patenting and product innovation is mainly driven by higher-quality granted patents. Likewise, Table A.V in the Appendix D provides similar results for other quality measures of patents (citations and claims), and Table A.IX in Appendix E shows that the correlation is weak for food-related categories and is mostly driven by products in categories such as health and beauty care, non-food grocery, and general merchandise.

The estimated coefficients capture the relationship between product introduction and patents associated with products. Not all patents, however, necessarily relate to product improvements: some patents may relate to cost savings from improvements to the firm’s general production processes. Nevertheless, our firm \times category data set filters out patents that are not specifically related to product introductions. Hence, to a large extent, our estimates should be driven by product patents rather than general process or organizational patents. We find a strong evidence supporting this point when we employ independent proxies for product-related and process-related patents drawn from claims texts as in Bena and Simintzi (2017). We find that the coefficient on product-related patents is essentially same as our benchmark coefficient, while process-related patents are entirely unrelated to measures of product innovation (Section A.II and Table A.VI in the Appendix).

Dynamics of the Effects – We are now interested in evaluating the timing of the effects captured in (2) and (3). Thus, we study the relationship between patents and product innovation by running the following separate linear regressions using the firm-category level data set:

$$Y_{ijt+k} = \beta_k E_{ijt} + \alpha_{ij} + \gamma_{jt} + u_{ijt+k} , \quad k = -4, \dots, 0, \dots, 4 \quad (4)$$

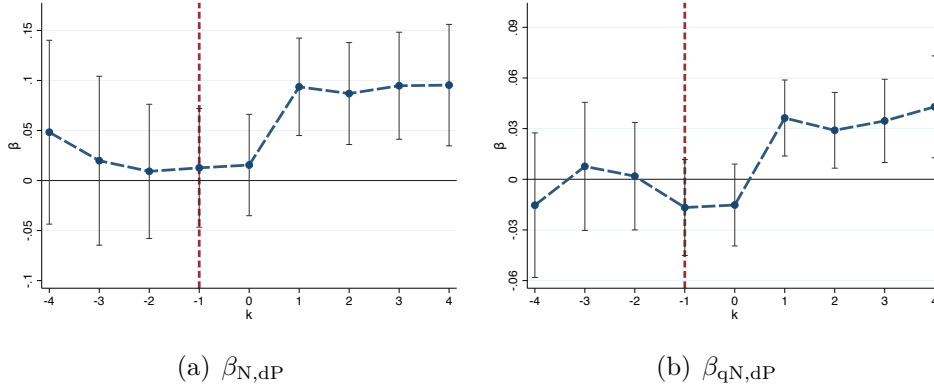
where Y_{ijt+k} is an outcome of firm i in product category j in $t+k$ associated with product introduction and E_{ijt} is either dP_{ijt} (as before, a dummy equal to one after firm starts patenting in category j) or $\log P_{ijt}$, which again denotes the log number of patents filed by firm i in product category j in t . We also include firm-product category and time-product category fixed effects.

Figure 2 plots the estimated coefficients β_k over k . The top panel shows the evolution of N and qN around the time at which the firm starts patenting in a certain product category. The bottom panel is about the intensive margin of patenting, and both are based on patent application years. Consistent with the results above, we find a positive association between patents and product introduction. Our estimates indicate that firms introduce about 10% more products after filing their first patent, with no pre-trends in outcomes before the firm switches to patenting (and 3-4% if we adjust for the novelty of new products– see (a) and (b)). The positive association reaches its maximum magnitude shortly after the first patent is filed in a product category and is fairly persistent thereafter.

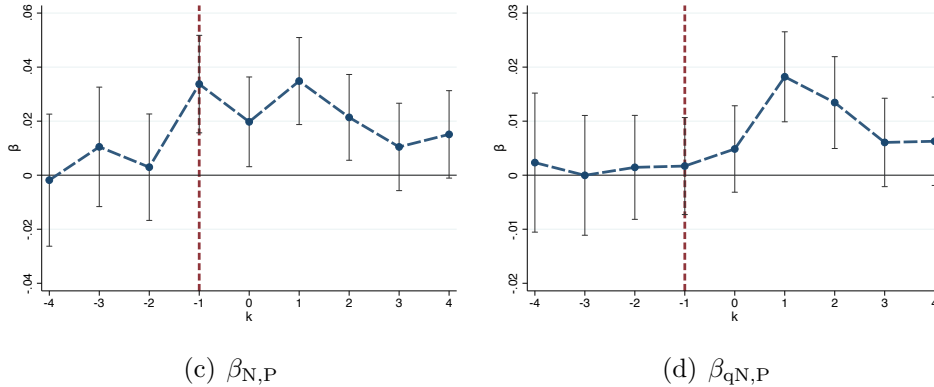
Likewise, our results (see (c) and (d)) exploring the co-movement between patent applications and product introduction indicate that product innovation spikes one year after new patent applications. With an exception for product introduction at $k = -1$ in (c), we do not find a significant relation for k below zero. These dynamic specifications are useful for inferring the long-run elasticity of product introduction to patents, in contrast to the instantaneous elasticities discussed previously. Unlike the results with first-time patent filers, the results for the intensive margin of patenting are not persistent over time, which indicates that filing an extra patent application is not associated with an incremental product introduction in the long run. Under exogeneity assumptions in the context of linear local projections (Jorda, 2005), the implicit long-run elasticity between patents and product introduction is the sum of the β_k coefficients from $k = 0$ onward. Our results point to an elasticity of about 0.1 for product introduction, and about 0.04 for quality-adjusted product introduction in the four years after the patent is filed. In Figure A.12 in the Appendix, we also show that other variables such as the stock of products or sales significantly increase after patents. We also confirm that our results are robust to considering the firm-level data from Match 1.

Figure 2: Product Innovation and Patenting: Dynamics

— *Switching to patenting* —



— *Number of patents* —



Note: The figure plots the estimated coefficients after estimating equation (4) for log product introduction, N , in (a) and (c), and quality-adjusted product introduction, qN , in (b) and (d). Our benchmark quality measure is defined in Section 3.1. The main explanatory variable in (a) and (b) is a dummy equal to one after the firm’s first patent in a product category and log number of patent applications in (c) and (d). The inverse hyperbolic sine transformation is used for logs. The vertical bands represent $\pm 1.65 \times$ st. error of each point estimate. Standard errors are clustered at the firm \times category level.

Taking Stock — The results so far are informative about our understanding of patent statistics as metrics of product innovation in the market. Our data set is uniquely suited to this analysis as we can directly associate specific new products with their underlying patents. We have somewhat contrasting findings. On the one hand, patents fail to capture a large fraction of actual innovation in the market. On the other hand, we find that patents are positively correlated with product innovation using within-firm and detailed product category variation over time. Our interpretation of the positive correlation between patents and product innovation is that, on average, firms come up with underlying ideas for new products and file patents to protect these ideas from being utilized by competitors; simultaneously, they develop these ideas into new consumer products. Our results support other papers’ findings that patenting is often associated with other real changes at the firm level,

such as increases in the firm’s scope or its stock market value (Hall et al., 2005; Balasubramanian and Sivadasan, 2011; Kogan et al., 2017).

A separate question is whether, conditional on a patent *application*, the *grant* of a patent may causally lead to future higher product innovation rates by firms. Patents approved by a Patent Office confer exclusion rights to the patent owner, and the benefits from these property rights may steer firms to introduce new products. For example, successful patents may be a positive signal for investors who might help alleviate credit constraints in the process of product development; patent holding may also boost consumers’ perception of how innovative the firm/product is and increase sales; affirmation of the property rights can also incentivize firms to enter the market where they can be assured that they would face temporarily low competition from imitators. Our results speak to the properties of patents as proxies for product innovation and not to the equally important but different question of whether a grant of patent rights casually leads to product innovations.²²

5 Product Innovation, Patents, and Competition: The Role of Firm Size

Our previous results show that, on average, patents carry a productive signal about actual product innovations in the market. But, in addition to reflecting forthcoming innovations, patents have a protective role by providing the legal right to exclude others from exploiting the same or similar inventions. As a result, firms can strategically use patents to defend their inventions, reduce competitive pressure, and deter entry (Cohen et al., 2000; Lanjouw and Schankerman, 2001; Jaffe and Lerner, 2004; Hall and Harhoff, 2012; Bloom et al., 2013). In this section, we study whether market leaders have stronger incentives to use patents for these strategic reasons by empirically evaluating how product innovation and patenting vary systematically with a firm’s market lead. We start by showing that:

Fact 3: Larger firms have lower product innovation rates (quantity and quality), but file more patents for each new product.

We then explore direct evidence on whether the protective role of patents in shielding firms from competition is larger among market leaders. We find that:

²²While not the main focus of our paper, we explore in Appendix G whether there is a causal link between patent grant and future product innovation. We follow the instrumental variables approach pioneered by Sampat and Williams (2019) and we find that, conditional on filing a patent, having a patent granted increases future product introduction.

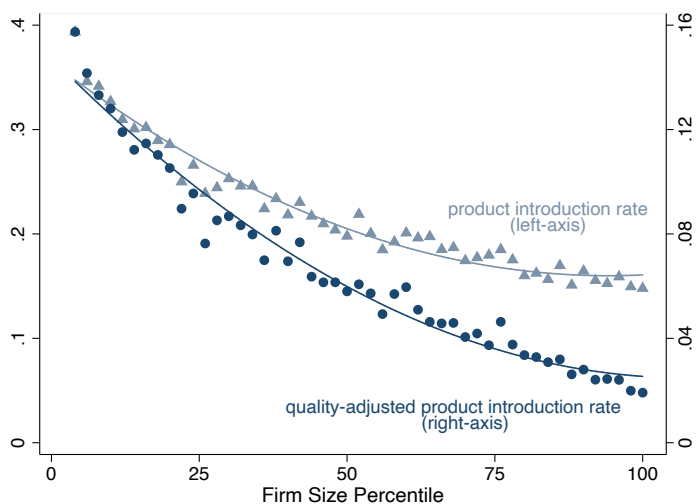
Fact 4: Patenting by larger firms is strongly associated with an increase in revenue above and beyond the patents' effect on product innovation.

Fact 5: Patenting by larger firms is associated with a decline in product introduction by competing firms.

5.1 Product Innovation and Patenting by Firm Size

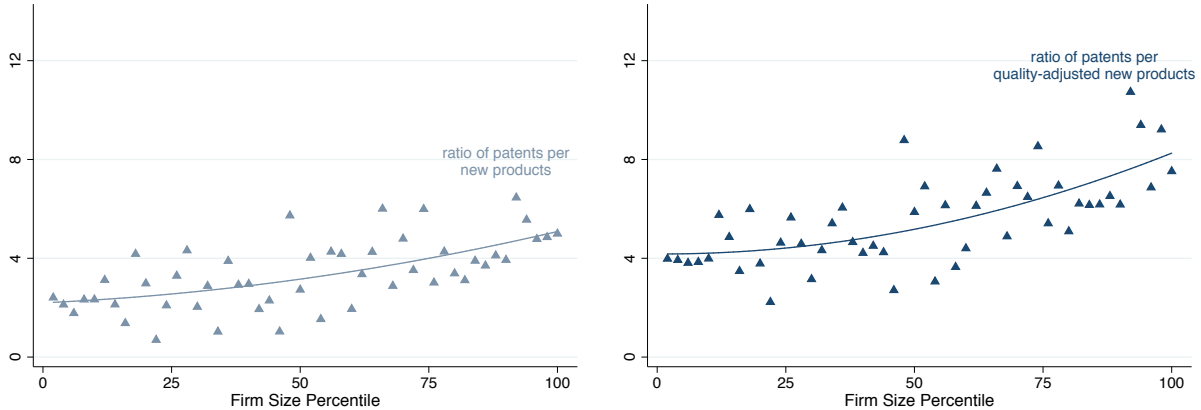
We begin by exploring how product innovation rates vary with firm size. Figure 3 plots the average product introduction rate – the ratio of product introduction to a firm's stock of existing products – for firms across product categories. Larger firms (within product categories) have lower product innovation rates. On average, firms in the top sales quintile have annual innovation rates of about 16%, while firms in the bottom quintile have rates twice as large. Larger firms do not compensate for this decline in the rate of new product introduction with innovations of higher quality. On average, firms in the top sales quintile have quality-adjusted product introduction rates of 3%, while firms in the bottom sales quintile have rates four times larger. The fact that the quality-adjusted introduction rate declines more steeply than the simple product introduction rate indicates that, on average, new products introduced by larger firms represent only incremental improvements over

Figure 3: Product Innovation Rate by Firm Size



Notes: This figure plots the relationship between product innovation rates and the relative size of the firm, defined by the firm's sales. We use the firm \times product category level data for the period 2007–2015, restricting the analysis to observations with sales above \$1,000. For each firm \times product category, we compute average sales, the average product innovation rate (new products divided by the total number of products sold), and the quality-adjusted product innovation rate (quality-adjusted new products divided by the total number of products sold). Within each product category, we assign firms to 50 bins for average sales and plot the average product innovation rate and the quality-adjusted product innovation rate for each bin. Each dot/triangle plots the averages after weighting each product category by its importance in the whole sector, as measured by the share of sales accounted for by the category.

Figure 4: Patents per New Products, by Size



Notes: This figure plots the relationship between the ratio of patent applications per new products and firm size as defined by sales. We use the firm \times product category level data set for the period 2007–2015, restricting the analysis to observations with sales above \$1,000. For each firm \times product category, we compute average sales, the average number of patent applications per new products, and the average number of patent applications per quality-adjusted new products. Within each product category, we assign firms to 50 bins of size based on average sales, and compute the average ratio of patents per new products and average ratio of patents per quality-adjusted new products for each bin. Each triangle plots the averages after weighting different product categories by their importance in the whole sector, as measured by their share of sales. The left figure plots the log ratio of patents per new products ($\times 1000$), and the right figure plots the log ratio of patents per quality-adjusted new products ($\times 1000$).

existing products and are thus less novel. Figure A.13 in the Appendix confirms similar patterns using alternative novelty metrics.

Next, we explore how patenting activity varies with firm size. With our data set, we can simultaneously measure both product innovation and the associated patent applications by firms. Figure 4 shows that larger firms, on average, file more patents for each new product introduced.²³ Note that this higher intensity of patenting activity relative to the number of new products introduced is not explained by the possibility that larger firms introduce fewer but more novel products: as one can see, after we adjust for the quality of new products, small and large firms’ innovation rates diverge even more.

We further explore how the relationship between product innovation and patents varies with firm size by estimating equation (3) for both product introduction, N , and quality-adjusted product introduction, qN , after controlling for time \times product category and firm \times product category fixed effects. By using within time \times product category and firm \times product category variation, we ensure that our results are not driven by potential confounders such as differences in patentability across firms and product categories. Table 6 reports the estimated coefficients for firms in different size groups. In line with the results discussed above, we estimate an average coefficient of 0.038 (column “All”). The table shows that

²³If we do not scale our measures of patenting, results are even starker: the unconditional probability of patenting and the total number of patents filed by large firms are much higher than they are for small firms (see Figure A.14 in the Appendix).

Table 6: Product Innovation and Patenting: by Size

| | Log N (t) | | | Log qN (t) | | |
|---------------|---------------------|---------------------|--------------------|---------------------|---------------------|---------------------|
| | All | Small | Large | All | Small | Large |
| Log $P(t-1)$ | 0.038*** (0.007) | 0.059*** (0.018) | 0.030** (0.013) | 0.019*** (0.003) | 0.033*** (0.007) | 0.017*** (0.006) |
| Observations | 409,641 | 61,350 | 86,953 | 409,641 | 61,350 | 86,953 |
| R-squared | 0.692 | 0.463 | 0.742 | 0.623 | 0.407 | 0.686 |
| Time-Category | Y | Y | Y | Y | Y | Y |
| Firm-Category | Y | Y | Y | Y | Y | Y |

Notes: The table shows regressions of the log number of new products ($\log qN$) and of log quality-adjusted new products ($\log qN$) in a firm \times category over time as a function of the log number of patents. P is the number of patent applications for a firm \times category \times year. For each firm \times product category, we define size based on the average sales over our sample period. The “All” column shows data for all sizes. “Small” column is restricted to the bottom size quintile. “Large” is restricted to the top size quintile. The inverse hyperbolic sine transformation is used for logarithms.

the relationship between patents and product innovation weakens with a firm’s size: larger firms in the top sales quintile have an elasticity twice as small as that of firms in the bottom sales quintile (0.030 versus 0.059). The results are similar for the quality-adjusted product introduction shown in the last three columns of the table. Also, the relationship between patents and product innovation weakens with firm’s size for both food and non-food categories (Figure A.16 in Appendix E), suggesting that the results are not driven by specific product categories and prevails across different types of sectors. Overall, while market leaders have the highest rates of patenting, the patents they file are less likely to translate into new products. This finding is by itself important from the perspective of measuring innovation in the market: for market leaders, patents may be misleading proxies for the actual innovation that drives productivity growth.

We perform a battery of robustness checks to confirm that our empirical findings are not driven by measurement issues. First, we examine whether the relatively weaker association between patent and product introduction of larger firms could be explained by differences in data coverage across firms of different sizes. The patent data set covers the entire portfolio of patents of firms but our product data set does not cover products outside the consumer goods sector, which could in turn result in lower rates of attribution of patents to new products for firms that also produce in other sectors. To allay this concern, we note that our empirical findings are based on our Match 2 algorithm, which filters out patents that are not related to the consumer goods sector. Moreover, we obtain similar results when we use a sample of firms that sell exclusively CPG products (see Section A.I for details on the construction of this sample).

Second, we also assess the possibility that our textual analysis of patents artificially

weakens the relationship between patents and products of large firms. The potential concern is that the text of patents filed by firms of different sizes may be systematically different and that our matching algorithm could be less effective in ascribing patents filed by larger firms to specific product categories. To better gauge this concern, we study textual characteristics of patents such as patent document length, number of unique words, textual diversity, and relative entropy of patents' word distribution, and we evaluate whether these characteristics vary systematically across firms of different sizes within the same product categories. We find no systematic differences in the textual characteristics of patents filed by large and small firms. Furthermore, we also do not observe significant differences in the share of matched patents and in the average similarity score across firm size (see Figure A.17 in Appendix for details). Overall, our exercises indicate that differences in data coverage and in the properties of the matching algorithm seem unlikely to explain the weaker association between patents and product innovation for large firms.

We also study whether the relatively weaker association between patents and product innovation of larger firms could be explained by economic factors, other than the strategic use of patents. One possibility is that larger firms shift toward process patents as they grow (Cohen and Klepper, 1996), thereby weakening the relationship between total patents and product introduction for larger firms. Using alternative proxies for product-related and process-related patents constructed in Appendix Section A.II, we find no systematic relationship between the share of this independent measure of process patents in the firms' portfolio and the firm's size (Figure A.18 in the Appendix). Moreover, if cost reductions due to process innovations are reflected in lower subsequent prices, we can test whether future price changes of larger firms react to patents more. However, we do not find such relationship in the data.²⁴

Lastly, there is no evidence that large firms do more experimental research and/or take more time to commercialize their inventions that could explain a weaker association between patents and products. Using the dynamic specifications of equation (3), we do not find evidence that patents held by larger firms are associated with product innovation with a longer delay.²⁵

²⁴A similar concern is that larger firms may file patents not to commercialize products but to license those patents to other firms. We assign patents to the patent holders and do not have information on temporary licensing agreements for all patents (such data do not exist). Moreover, prior work (Fosfuri, 2004; Gambardella et al., 2006), suggests that, if anything, larger firms are less likely to license their patents out.

²⁵Indeed, recall that in Section 4, we saw that the increase in product innovation after a new patent filing was short-term.

5.2 The Role of Patents for Competition by Firm Size

Our previous finding that the association between patents and product innovation weakens for larger firms is suggestive that market leaders use patents as a strategic tool to protect their market shares. Firms may advance their market position not just by introducing new and innovative products, but by accumulating patents that reduce competition and help the firm fend off threats that competitors pose.²⁶ We take advantage of the richness of our data set to provide direct evidence on this mechanism.

First, we use information on product sales and prices to quantify the role of patents for firms of various sizes. We use the following specification:

$$\Delta \log \text{Sales}_{ijt} = \psi \log P_{ijt-1} + \rho \log N_{ijt} + \theta_{ij} + \gamma_{jt} + \varepsilon_{ijt} \quad (5)$$

where the dependent variable is the logarithm of the change in sales at time t , $\log P_{ijt-1}$ is the total number of patent applications until time $t - 1$, and $\log N_{ijt}$ is the number of new products introduced at t (we also use the quality-adjusted product introduction $\log qN_{ijt}$). Our coefficient of interest is ψ , which measures the relationship between sales growth and patents after accounting for the effect that patents may have on sales through increased product innovation.

Table 7 shows the results for all firms and for firms grouped according to size. Overall, we find a positive significant relationship between patents and future growth in sales even after controlling for product innovation (columns “All”). This finding suggests that holding an additional patent allows firms to increase its sales even after accounting for the increase in sales that results from new product offerings. Importantly, this effect is highly heterogeneous across firm size. For firms in the bottom sales quintile (columns “Small”), there is no statistical association between patents and sales growth after we control for product introduction. However, for firms in the top quintile (columns “Large”), we find that an increase in total patent applications has a significant positive association with sales growth above and beyond its effect through product introduction. Note also that the direct impact of product innovation on sales growth (coefficients on $\log N(t)$ and $\log qN(t)$) decreases as firms increase in size. Hence, by splitting the sample into small and large firms, we learn that, while both patents and new products are associated with increased future sales, the conditional impact of new products is more important for smaller firms, while the impact of patents is important for larger firms.

²⁶The accumulation of patents often creates a web of overlapping intellectual rights which make it difficult for competitors to approach the market leader’s technology domain and to leapfrog them. See, for example, [Shapiro \(2000\)](#) for the discussion of patent thickets.

Table 7: Patenting and Sales Growth

| | $\Delta \text{ Log Sales (t)}$ | | | $\Delta \text{ Log Sales (t)}$ | | |
|---------------|--------------------------------|---------------------|---------------------|--------------------------------|---------------------|---------------------|
| | All | Small | Large | All | Small | Large |
| Log P(t-1) | 0.061*** (0.016) | -0.081 (0.077) | 0.099*** (0.019) | 0.073*** (0.016) | -0.101 (0.077) | 0.111*** (0.019) |
| Log N(t) | 0.265*** (0.003) | 0.316*** (0.011) | 0.160*** (0.004) | | | |
| Log qN(t) | | | | 0.406*** (0.006) | 0.581*** (0.029) | 0.215*** (0.007) |
| Observations | 296,320 | 40,666 | 65,680 | 296,320 | 40,666 | 65,680 |
| R-squared | 0.291 | 0.377 | 0.296 | 0.275 | 0.368 | 0.281 |
| Time-Category | Y | Y | Y | Y | Y | Y |
| Firm-Category | Y | Y | Y | Y | Y | Y |

Notes: The table presents estimated outcomes of changes in log sales at the firm \times category level as a function of the log number of patent applications by until time $t - 1$ and the log number of new products introduced at time t (or quality adjusted new products), by size groups. We use the firm \times product category data set for the period 2007–2015, restricting the analysis to observations with sales above \$1,000. For each firm \times product category, we define size based on average sales over the sample period. “All” column uses data for all sizes. “Small” column is restricted to the lowest size quintile. “Large” is restricted to the top size quintile. The inverse hyperbolic sine transformation is used for logarithms.

We further explore if this difference in the association between sales and patenting between small and large firms operates through changes in prices and/or change in quantities sold (Table A.VII in Appendix shows the results using specification (5) for prices and quantities). We find that the additional sales premium of patents conditional on product innovation for large firms occurs both because of higher quantities sold and higher prices.

Our main conjecture is that this additional sales premium associated with patents of large firms is driven by the effects of these patents in staving off competition: if patents discourage competitors from introducing new products, patent holders will benefit by serving a larger market and charging higher prices. Next, we directly investigate whether patents by market leaders are associated with declining product introduction on the part of their competitors, who we will refer to, for simplicity, as market followers. We identify the market leader in each category as the firm with the highest sales in that category and the followers as the remaining firms operating in that market.²⁷ Then for each year t and market j , we compute the total number of new products introduced by the leader N_{jt}^L and by its followers N_{jt}^F in t , and we compute the total numbers of patent applications introduced by the leader P_{jt}^L and by its followers P_{jt}^F until t . We evaluate how product innovation by followers responds to

²⁷To have a static firm-level measure, we define leaders as of 2006, which is the first year of our data. However, the results are not sensitive to a different choice, like using average sales over all years. Moreover, we consider alternative definitions of market leaders (e.g. top decile) and the results are robust.

patenting (and product innovation) of the leaders as follows:

$$\log N_{jt}^F = \eta^F \log P_{jt-1}^L + \alpha^F \log N_{jt-1}^L + \theta_j^F + \gamma_t^F + \varepsilon_{jt}^F, \quad (6)$$

where η^F is our coefficient of interest, measuring the association of patents of leaders with the product introduction by followers. We control for $\ln N_{jt-1}^L$ to ensure that the relationship between leaders' patents and followers' product introduction is not driven by possible direct interactions between the leader's and followers' product offerings (such as learning from new products on the market).²⁸ We also include both time- and product-category-fixed effects to control for time trends and differences in the intensities of patenting and product innovation across product categories. Likewise, we estimate a symmetric regression that measures how leaders' innovation is affected by followers' patenting:

$$\log N_{jt}^L = \eta^L \log P_{jt-1}^F + \alpha^L \log N_{jt-1}^F + \theta_j^L + \gamma_t^L + \varepsilon_{jt}^L \quad (7)$$

These regressions help us test if the relation between patents of competitors and product introduction is affected by whether we focus on leaders or followers.

Table 8 presents the estimated coefficients. Column 1 shows that product introduction by followers is negatively correlated with the size of the leader's patent portfolio. This result suggests that followers reduce the introduction of new products in categories where the leader intensifies its patenting efforts. In column 2, we also control for total sales of the market to account for potential shifts over time in the importance of different types of products. In turn, columns 3 and 4 show that product innovation by leaders is not related to the followers' patenting activity. Hence, while patents can be thought of as a protective tool used to hinder competition in the product market, our results indicate that this hypothesis is likely to apply when patents are in the hands of large market leaders.

Having established that the weak relationship between patents and products for market leaders is largely explained by their use of patents for protective reasons, we next evaluate what are the characteristics of these patents and how they differ across firm's size. While patents on new technologies should generally deter competitors from imitating those technologies, certain strategic patents go beyond this goal and discourage entry of products even if these products are only loosely related to the patented product. A good example is the patent for P&G Swiffer Wet Jet mops. Instead of patenting the features of the invention, P&G patented the specific functionality of the disposable cloths. The original patent and the more than 80 follow-up patents have made it difficult for competitors to enter the

²⁸We also use quality-adjusted new products in all of these regressions, and the results are similar.

Table 8: Patenting of Market Leaders and Followers

| | Followers Log N^F | | Leaders Log N^L | |
|-----------------|------------------------|----------------------|----------------------|----------------------------------------|
| | (1) | (2) | (3) | (4) |
| Leaders | | | Followers | |
| Log P^L (t-1) | -0.071*** (0.007) | -0.059*** (0.007) | Log P^F (t-1) | -0.015 (0.047) -0.012 (0.044) |
| Log N^L (t-1) | 0.010*** (0.002) | 0.005* (0.002) | Log N^F (t-1) | 0.215* (0.112) 0.185* (0.094) |
| Observations | 3,192 | 3,192 | Observations | 3,188 3,188 |
| Category | Y | Y | Category | Y Y |
| Time | Y | Y | Time | Y Y |
| Controls | N | Y | Controls | N Y |

Notes: The table shows the relationship between the patents of leaders (followers) and the product introduction of followers (leaders). The leader is defined as the firm with the highest sales in a given category in 2006; the followers are defined as the rest of the firms in the categories. In columns (1) and (2), the dependent variable is the log number of products introduced by followers at time t , and the independent variables are the log number of patent applications by leaders until time $t - 1$ and the log number of new products introduced by the leader at time $t - 1$. In columns (3) and (4), the dependent variable is the log number of products introduced by leaders at time t , and the independent variables are the log number of patent applications filed by followers until time $t - 1$ and the log number of new products introduced by the followers at time $t - 1$. Columns (2) and (4) also control for total sales in the category-time. The inverse hyperbolic sine transformation is used for logarithms.

market. Indeed, during our sample period, generic sweeper mops were basically absent. P&G has a market share of approximately 95% in sweeper mops – much larger than that of leaders in other categories whose share is 40-50% on average. It is also not unusual to find examples of firms with “sleeping” patents that deter entry but do not lead to products on the market. An example is Driscoll’s, which controls a third of the U.S. berry market. The company invests heavily in a breeding program to develop new berry varieties which they patent but often do not commercialize. Driscoll’s has one of the highest ratios of patents per new product in our data. The company has also recently been involved in several lawsuits to protect its patent portfolio from potential competitors.²⁹

Although we cannot identify every specific patent that does not lead to product innovation but hinders competition, we can nevertheless identify some general patent characteristics by comparing patents of market leaders to those of the followers along various dimensions, such as their novelty and impact. We find that market leaders’ patents accumulate fewer forward citations hence leading to lower follow-up research, have a larger share of self-citations, and exhibit higher textual similarity with respect to preceding patents. We also find that these patents have broader patent claims and are twice more likely to be engaged in litigation (see Table A.VIII in the Appendix for more details on this analysis).

We view the results laid out in this section as providing a consistent story for the firms’

²⁹ “How Driscoll’s Reinvented the Strawberry”, The New Yorker, August 21, 2017).

use of various growth strategies. Firms may use both productive and protective strategies to grow their market shares. We document that as firms grow, they rely relatively less on productive strategies that encourage the introduction of new and improved products in the market and increasingly rely on protective strategies such as patenting. These patents often discourage competitors from introducing new products, allowing large firms to serve a larger market and charge higher prices.

6 Conceptual Framework

In this section, we offer a simple theoretical framework that illustrates the relationship between innovation, patenting, and creative destruction. Our goal is twofold. First, the framework is meant to build intuition about the incentives for patenting and product innovation for firms with different market positions, consistent with empirical patterns documented in the previous sections. Second, we use the model to perform a simple back-of-the-envelope calculation of the private value of a patented innovation, and we decompose this private value into its protective versus productive components. We define the productive component as the option value of implementing the patented idea into higher-quality products in the market, thereby increasing profits. The protective component is the value that the firm gains by impeding creative destruction.

Our framework builds on the quality-ladder model of innovation with creative destruction (e.g. [Aghion and Howitt, 1992](#)). In the model, product innovation takes the form of upgrades to the quality of products on the market. These innovations come from either the incumbent leader trying to prolong its lead or from market entrants aiming to become the new leaders. We consider an exercise in which the incumbent firm obtains an idea/blueprint for an innovation and makes a once-in-a-lifetime decision about commercializing that idea into a product and/or patenting it. If the firm decides to commercialize the idea, it will gain additional profits when it introduces higher-quality products to the market; and profits exhibit decreasing returns in quality. In turn, patenting the idea grants the firm extra protection against creative destruction from entrants. Both patenting and product innovation are costly activities.

The key result of our analysis is that product innovation and patenting decisions vary with the incumbent size, which maps to our empirical result that firms of different sizes have different propensities to innovate and patent. In the model, for the same idea/blueprint, smaller firms decide to commercialize ideas into better-quality products, mid-size firms do both product innovation and patenting, and very large firms file patents protecting their market shares but do not upgrade their products on the market.

Production – Consider a partial equilibrium framework that describes innovation in a single product category. Given that our main empirical facts use the variation within product categories, our focus is on single-category firms and their incentives for innovation and patenting in that category.³⁰ There are M potential producers, and aggregate output is a combination of quality-weighted varieties:

$$Y = \frac{1}{1-\beta} \left[\sum_{m=1}^M q_m^{\frac{\alpha}{1-\beta}} y_m \right]^{1-\beta}, \quad 0 < \alpha < \beta < 1 \quad (8)$$

where y_m denotes the quantity and q_m is the quality level of variety m . This specification implies that products from different producers are perfect substitutes after adjusting for their qualities. The parameter α captures the consumer’s satiation with respect to additional quality. Labor is the only factor of production. Producers use labor to produce output by hiring labor at the common wage rate of w . Output of variety m is then given by $y_m = l_m$, where l_m is labor used to produce variety m . We assume that the overhead cost of production ϵ must be paid before choosing prices and output. Since producers’ marginal costs are the same and qualities are different, under Bertrand competition, this overhead cost allows the highest-quality firm to win the market and act as a single producer.³¹

The producer maximizes profits by choosing the price of its product subject to demand from (8),³² which delivers the following equilibrium objects for output (y), sales (R), and profits (Π), respectively (hereafter, we drop the subscript m):

$$y = \frac{1-\beta}{\beta} \frac{\pi}{w} q^\gamma, \quad R = \frac{\pi}{\beta} q^\gamma, \quad \Pi = \pi q^\gamma, \quad (9)$$

where $\pi \equiv \beta \left(\frac{1-\beta}{w}\right)^{\frac{1-\beta}{\beta}}$ and $\gamma \equiv \frac{\alpha}{\beta}$. Hence, firms with higher-quality products are larger, and generate higher sales and profits. Moreover, since $\gamma < 1$ because $0 < \alpha < \beta < 1$, the marginal quantity, sales, and profits decrease with quality.

Product innovation and patenting choices – Now consider the once-in-a-lifetime decision of product quality upgrade and patenting for an incumbent with quality q who

³⁰All our subsequent qualitative results can be generalized to the case of multi-category firms at the expense of the model tractability. In fact, the existence of multiple product categories with potential patent protection spillovers between them will amplify our results on firms’ patenting incentives.

³¹This assumption simplifies the setup. Alternatively, we would need to work with limit pricing, where the firm with highest quality would still capture the entire market, but the price would be determined by the price of the second highest quality producer.

³²Price of Y is normalized to one.

exogenously obtains an idea of size λ .³³ We think of q as the aggregate “composite” quality level of various products that firms currently offer in particular product category, and the parameter λ represents the composite measure of product innovation we defined in Section 3. If the firm decides to upgrade the quality of its products from q to $q + \lambda$ to generate higher profits, it also has to pay the costs of product development and commercialization c_m . In this case, $q + \lambda$ becomes the largest available quality in the economy. Simultaneously, the firm can also decide to patent the blueprint at a cost c_p .³⁴ A patent grants the firm additional protection against being replaced by an entrant (more details below). If the firm decides to patent, even if the idea is not commercialized, the highest-available quality in the economy becomes $q + \lambda$ since, by patenting, the firm makes the idea “public”. Note that the highest-available quality in the economy could be different than that commercialized by the firm and available to the consumers. In this sense, firms’ activities in the product and patent spaces are separated. Product innovation does not necessarily imply patenting activity, and neither does introducing a patent necessarily imply product innovation.

Creative destruction via entrants – Incumbents can be replaced by entrants through creative destruction. The model includes an exogenous arrival rate of entrants at each instant p . Entrants build on “the shoulders of giants” and can replace incumbents by improving upon the highest-quality product available in the economy. The underlying assumption is that entrants can learn from products available in the market and from patents. Hence, “the shoulders of giants” correspond to $q + \lambda$ unless the incumbent neither upgrades nor patents, in which case the highest available quality is q . Entrants draw innovation of step size λ^ϵ from a uniform distribution on $(0, 1)$. Patenting protects the quality level of incumbents $q + \lambda$ by creating a “wall” of height $\epsilon > 0$ that entrants need to overcome to enter the market. The parameter ϵ captures the condition that entrants need to come up with an innovation sufficiently different from what has been patented before, which can depend on the strength of intellectual property protections as well as the scope of the patent. Given these assumptions, the probability of creative destruction is p if the incumbent does not patent and is $p(1 - \epsilon)$ if the incumbent patents (Appendix H provides the proof). Notice that, in contrast to standard models of creative destruction, not all product quality improvements by entrants will find their way to the market. The separation between the patent space and the product space introduces the possibility that a better-quality product

³³For simplicity, we assume that this is a one-time choice. Hence, the idea is either used or disappears afterwards. A more-complete approach with a dynamic decision of patenting and product innovation would bring similar tradeoffs at the expense of tracking the evolution of a firm’s position both in the product and patent spaces.

³⁴We think of c_p as the combination of research, legal filing, and potential patent enforcement costs.

is not introduced to the market because it is blocked by existing patents.

Value functions – Let us denote the (gross) value of a firm with existing product quality q that both upgrades quality and files patents as V^{11} , the value of only upgrading as V^{10} , the value of only patenting as V^{01} , and the value neither upgrading nor patenting as V^{00} :

$$\begin{aligned} V^{11}(q) &= \frac{\pi(q + \lambda)^\gamma}{r + p(1 - \varepsilon)}, & V^{10}(q) &= \frac{\pi(q + \lambda)^\gamma}{r + p}, \\ V^{01}(q) &= \frac{\pi q^\gamma}{r + p(1 - \varepsilon)}, & V^{00}(q) &= \frac{\pi q^\gamma}{r + p}. \end{aligned}$$

Then, the value function of the incumbent firm with existing quality q is

$$V(q) = \max \left\{ V^{11}(q) - c_m - c_p, V^{10}(q) - c_m, V^{01}(q) - c_p, V^{00}(q) \right\}. \quad (10)$$

Comparative statics and empirical facts – To simplify the setup, we rely on standard assumptions in this class of models and analyze a problem of a single producer in a product category. Two clarifications on how we match up the model to the data are in order. First, in what follows, our exercise is to do comparative statics of product innovation and patenting decisions of firms of different sizes (sales or profits), here captured by q . Second, the intuition we build from this comparative statics can be extended to a similar setting with multiple firms holding different market shares – alas, at the expense of tractability of an illustrative model.

Turning to comparative statics, we notice that incentives for product innovation decline as firm size increases, while the returns to patenting increase. Because marginal profits decrease as q increases, the incremental returns from product innovation decline with firm size, which is the same intuition that underlies the well-known *Arrow-replacement effect*.³⁵ Our specification captures how larger firms find it less profitable to replace themselves due to larger cannibalization effects. On the other hand, the returns to patenting increase with size as larger firms have a higher value to protect. In fact, we show in Appendix H that under mild conditions on costs, in this economy there exist cutoffs q^* and q^{**} such that a firm upgrades without patenting when $q < q^*$, engages in both product innovation and patenting in the intermediate region $q^* < q < q^{**}$, and only patents when $q > q^{**}$.³⁶

³⁵We provide an empirical estimate of γ and confirm that it is lower than one. However, instead of these decreasing returns that generate a declining relationship between size and innovation, one can generate it through other ways, such as by implementing weaker scalability of R&D technology with increasing size as in Akcigit and Kerr (2018) or an innovation-advertising trade-off as in Cavenaile and Roldan (2020).

³⁶The required conditions on the costs c_m and c_p ensure that at least one firm finds it profitable to introduce a product and at least one firm finds it too costly to patent. This sharp cutoff rule clearly hinges

Overall, the model delivers an equilibrium that rationalizes the main empirical patterns uncovered in the previous sections. First, *Fact 1* indicates that many firms innovate without patenting. Through the lens of our model, firms below the cutoff q^* will make such a decision. The likelihood that firms innovate without patenting increases with the cost of patent filing c_p and declines with the incremental benefits associated with deterring competition ε . Second, our model can generate a weak positive relationship between patent filing and subsequent product innovation within firms, consistent with *Fact 2*. The strength of this relationship in the model depends on the relative shares of firms in the $[q^*, q^{**}]$ and $[q^{**}, \infty)$ regions. Third, as explained above, the model also rationalizes the main empirical heterogeneity in the data outlined in *Fact 3* – product innovation declines and patent filing increases with firm size, q . This implication also explains the weak correlation between patents and subsequent product innovation for largest firms shown in Table 6. The model also speaks to our empirical *Fact 5* on the relationship between patenting, firm size, and competition. By construction, patents in the model reduce creative destruction. At the same time, the model implies that larger firms rely on patenting more. Hence, larger firms also face a lower risk of creative destruction. Finally, in the next section, we will see how the model speaks to our last fact on how the returns to patenting vary by firm size.

Decomposing the value of a patent – We now use the model to evaluate the (gross) *private* value of a patent to a firm. We define the value of a patent as the incremental gain in firm value that a patented innovation provides: $V^{11}(q) - V^{00}(q)$.³⁷ These are incremental discounted profits that come both from a blueprint/idea contained in a patent and also from the legal protection for that idea. We use the model to differentiate between these two components embedded in the private patent value – productive and protective values.³⁸ Productive value comes from the option value of commercializing an idea. In our model,

on the simplifying assumption of fixed innovation and patenting costs that buys us tractability. However, the main features of the model can be easily generalized with a more realistic cost structure. For example, if larger firms are more experienced in patent filings and have bigger resources to engage in litigation, this would increase even further their patenting incentives. On the other hand, large litigation costs and financial constraints could hold up smaller firms from patenting more than in the current model.

³⁷Note that these values are gross of patenting and commercialization costs. When comparing these size-dependent values to c_l and c_p costs, firms’ optimality results in the innovation and patenting thresholds discussed above.

³⁸To relate this definition to the existing approaches studying patent value, consider the following examples. One can evaluate private patent value by estimating a stock market response to patent filings (or grants) (Hall et al., 2007; Kogan et al., 2017) or by looking at patent sales data (Abrams et al., 2013). In the former case, to the investors, patent filing indicates both arrival of new idea that could materialize into innovations as well as a firm’s intent to protect the existing or future rents from competitors’ entry. In the latter case, when firms sell patents, they sell both idea/blueprint as well as the right to protect it. Using our model, we are able to explicitly differentiate between these two distinct components of the patent value.

this is the value of implementing the innovation blueprint of size λ . In turn, protective value comes from the ability of a patent to protect firms' market lead from competitors. In our model, this is the value of exercising patent protection of size ε . Hence, we obtain

$$\begin{aligned} \text{Patent Value}(q) &= V^{11}(q) - V^{00}(q) & (11) \\ &= \underbrace{\frac{\pi(q + \lambda)^\gamma}{r + p(1 - \varepsilon)} - \frac{\pi(q + \lambda)^\gamma}{r + p}}_{\text{Protective}} + \underbrace{\frac{\pi(q + \lambda)^\gamma}{r + p} - \frac{\pi q^\gamma}{r + p}}_{\text{Productive}} \end{aligned}$$

The second line decomposes the total patent value into productive and protective components by adding and subtracting $V^{10}(q)$. Productive value denotes the profit premium from commercializing a product of upgraded quality if we hold creative destruction fixed, while protective value denotes the profit premium from lower creative destruction holding the technology of a firm fixed. Notice that since the values depend on q , this implies that the same patent (with same idea of size λ and protection of size ε) is valued differently by firms depending on their existing size. Specifically, productive value, or value of product innovation, declines as firms grow: the same amount of innovation brings marginally lower returns. In contrast, protective value increases as firms grow: the use of patent protection is more beneficial as the value of the firm increases. This observation speaks to empirical *Fact 4* documenting the heterogeneity in the revenue premium from a patent by firm size. Consistent with Table 7, in the model, the return to a patented innovation comes both from product innovation and protection; and the latter is more important for larger firms.³⁹

We now set the parameters of the model to estimate the average value of a patent for firms in our data. To estimate (11), we need to assign values to π , λ , γ , p , and ε . First, we normalize the average quality within each product category in our data to one. Notice that we do not observe profits, but given (9), we know that sales are proportional to profits such that $\Pi = \frac{\mu-1}{\mu} \times R$, where μ is the markup. The profit of an average firm is then $\frac{\mu-1}{\mu}$ multiplied by the sales of the average firm, which we take to be equal to the average yearly sales of all firms across all product categories (1.36 million 2015 USD). We take $\mu = 1.21$ drawing upon Barkai (2020)'s average estimate of markups in the U.S. economy in 2014. Assigning values to p and $p(1 - \varepsilon)$ involves the following considerations. In the model, if firms do not innovate they face a creative destruction rate that leads to the decline of their expected sales. Hence, we infer the values of p from firms' growth in sales when they do

³⁹In the data we do not observe profits and use revenue for our measure of returns, while (11) is expressed in terms of profits. To connect these two measures of returns, notice that in the model revenue and profits are proportional as seen from (9).

not introduce new products in a given year. In our data, the median firm that does not hold any patents suffers a loss in sales when it does not introduce new products: log sales change is equal to -10.3% . This decline in sales is attenuated if a firm holds a patent, thus giving an estimate for ε . The implied values for creative destruction are $p = 0.098$ and $p(1 - \varepsilon) = 0.095$.

Lastly, we jointly estimate λ and γ . Intuitively, λ determines the average growth when the firm innovates, and γ affects how this growth varies with firm size. Specifically, the model implies the following relationship between firm growth and relative size, conditional on product innovation:

$$\Delta \ln R_t = \gamma \ln \left(1 + \lambda \left[\frac{R_{t-1}}{\bar{R}_{t-1}} \right]^{-\frac{1}{\gamma}} \right)$$

We estimate this relationship with a non-linear least squares regression applied to the sample of firms who introduce new products in that year. We define the relative size of firms as sales divided by the average sales of firms in that year and product category. The resulting estimates are $\gamma = 0.899$ (s.e. 0.364) and $\lambda = 0.024$ (s.e. 0.008).

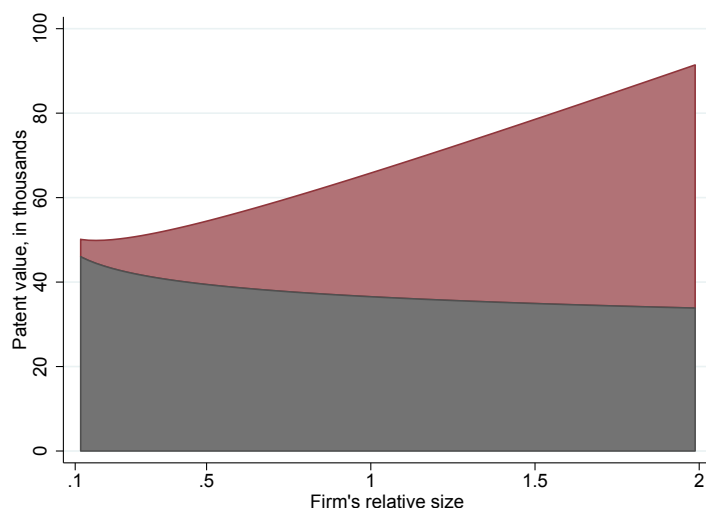
Figure 5 plots the value of a patent against the relative size of firms. The red shaded area depicts the contribution of the protective component of the patent's value, and the gray area depicts the contribution of the value's productive component. For the average firm the value of a patent is around \$65,000.⁴⁰ The estimated value increases drastically as firm size increases, mainly due to the contribution of the protective value. For example, for firms ten times smaller than the average firm, only 9% of the value comes from the protective component, while for firms that are twice as large as the average firm, the protective component accounts for 60% of a patent's value.

Our methodology for estimating the value of the patent differs greatly from those used previously in the literature as it relies on the structure of our model, the matched data we constructed between patents and products, and the realized sales of products observed in the data.⁴¹ Nonetheless, we find that our estimates are well in the range of other estimates reported in the literature. Using patent renewal information to infer the private value of U.S. patents issued in 1991, [Bessen \(2008\)](#) estimates a patent's mean value to be \$121,000 (median \$11,000). Interestingly, consistent with our results, [Bessen \(2008\)](#) also finds that the value of patents held by smaller firms is lower, while litigated patents are more valuable. [Serrano \(2010\)](#) estimates the average private value of holding a patent to be

⁴⁰Notice that our calculations do not include sales from the stores not covered by Nielsen. To get at the nationwide sales, we can roughly scale our sales twice (see Appendix A.I for details).

⁴¹The literature has used various methods to estimate the value of a patent: direct survey questions, inference from observed patent renewals by firms, stock market responses to patent news, as well as direct estimates from patent sales samples. For a comprehensive review, see [Hall and Harhoff \(2012\)](#).

Figure 5: Estimated Gross Private Patent Value



\$90,799 (median \$19,184). Using data from a large non-practicing entity, which presumably holds mostly valuable patents, [Abrams et al. \(2013\)](#) find that the mean value of a patent is \$235,723 (median \$47,955). The advantage of our methodology is that it allows us to decompose the patent value into its two inherent components – productive and protective. The decomposition uncovers the dual role of patenting and the role each component plays for firms of different sizes.

Overall, our theoretical framework illustrates how the incentives for patenting and product innovation go in opposite directions as firms grow. Because of cannibalization of their existing rents, larger firms give up on many ideas which smaller firms would find worth commercializing. Meanwhile, the incentives for patenting go in the opposite direction – larger firms want to protect their existing rents relatively more.

7 Conclusion

Using textual analysis of patent documents and product descriptions, we construct a new patent-to-products data set to study the relationship between patents and product innovation. We find that more than half of the product innovation is not associated with patents. Nonetheless, patent filing is positively associated with subsequent product innovation by firms, on average. We document substantial heterogeneity in this relationship. Patents filed by larger firms reflect less the actual product innovation than other patents do. Instead, we find strong evidence suggesting that the important role of patents for market leaders

is to deter product innovation of competitors and protect sales of their existing products. Hence, our results indicate that although on average patents capture product innovation in the market, because the relationship between patents and innovation changes with firm size, patent-based measures distort the differences in actual innovation between firms of different sizes.

Using a simple theoretical framework, we show that for the same patented idea, a larger firm can reap a greater incremental monetary return than a smaller firm can. However, for these large firms, more of this return is derived from the patent's ability to hinder competition than is derived from commercializing a new product using the patented idea. We argue that understanding the contribution of the productive and protective components of patenting and how they vary by firm size has important implications for our understanding of growth, innovation, and intellectual property policy. Specifically, policymakers should pay more attention to the state-dependence of patent and R&D-based policies; these policies should acknowledge that firms' incentives to use intellectual property vary greatly with firm size and market leadership.

In their comprehensive analysis of patent reform in the U.S., [Jaffe and Lerner \(2004\)](#) argue that the seemingly innocent changes in patent policies in the early 1980s significantly affected firms' incentives toward strategic patent filing. Understanding how the increasing reliance on strategic patenting has contributed to the recent trends of increasing dominance of large firms and declining business dynamism ([Decker et al., 2016](#)) is an interesting research agenda going forward.

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